

Executive Compensation and Firm Leverage

by

Michael Albert

Business Administration
Duke University

Date: _____

Approved:

David Robinson, Supervisor

John Graham

Lukas Schmid

Giuseppe Lopomo

Dissertation submitted in partial fulfillment of the requirements for the degree of
Doctor of Philosophy in Business Administration
in the Graduate School of Duke University
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ABSTRACT

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Abstract

This dissertation explores the role of executive compensation in determining the capital structure decisions of a firm. CEOs experience a large personal cost of default that interacts through the risk adjusted probability of default with their compensation contract. Since default happens in a particularly costly state of the world for a CEO whose compensation contract consists primarily of pay for performance elements, i.e. a CEO who has a large personal equity stake in the firm, a large pay performance sensitivity is negatively and significantly associated with firm leverage choice. I document this effect in detail for the first time, and I show that it is both statistically robust and significant in magnitude, approximately 1% of firm value. I show that this effect is driven by the stock holdings of the CEO, not the option holdings. I provide a simple principal agent model that explains the observed negative relationship and makes additional predictions on the relationship of other firm characteristics to pay performance sensitivity and leverage. I then test and confirm these predictions empirically using a standard OLS framework and an instrumental variable approach to control for endogeneity in the compensation contract. I also look at leverage adjustment speeds and show that CEOs with higher pay performance sensitivity adjust leverage upwards towards target values more slowly and downwards more quickly than their peers, and I interpret this as direct evidence that CEOs are actively managing personal risk through firm leverage choice.

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1

Introduction

Standard theories of corporate leverage assume that the observed leverage choice is optimal for maximizing the value of the firm. However, there is a large and growing literature that suggests that many firm choices are driven not by considerations of firm value maximization, but instead by the idiosyncratic effect of CEOs (e.g. Bertrand and Schoar (2003), Cadenillas et al. (2004), and Malmendier and Tate (2005)). In addition, there is a growing literature examining the effect of executive compensation on firm financing decisions (e.g. Brockman et al. (2010)). The importance of understanding the effect of the structure of managerial compensation on firm choices, specifically pay for performance components, has increased since the manner in which executive compensation is structured has changed dramatically over the past three decades. The median exposure of executive compensation to stock price tripled from 1984 to 1994 (Hall and Liebman (1998)) and further doubled between 1994 and 2000 (Bergstressor and Phillippon (2006)). Further, CEOs face significant negative shocks to lifetime income after being the active manager for a firm that experiences default (Eckbo and Thorburn (2003), Eckbo et al. (2012)). While CEO fixed effects have been shown to affect the leverage chosen by a firm (Frank and Goyal (2007)),

the question of how and why executive compensation affects the leverage choice of a firm is still open. This dissertation shows that due to the interaction of standard compensation contracts and personal default risk, pay performance sensitivity (hereafter referred to alternately as PPS or pay performance sensitivity) is significantly negatively related to firm leverage decisions through the risk adjusted probability of default increasing as more of a CEO's personal wealth portfolio is exposed to firm specific risk.

The effects of executive compensation on firm choice has been studied in other contexts. Aggarwal and Samwick (2006) show that firm performance increases in incentives. Likewise, Aggarwal and Samwick (1999) show that incentives influence firm pricing policy. Denis et al. (1997) show that CEO incentives affect the diversification decision of a firm. In order for executive compensation to play a role in determining the firm's capital structure three things must be true. First, the executive must have the ability to determine capital structure. Second, the executive must have an incentive to deviate from the interests of shareholders. Third, the compensation contract must affect the incentives of the executive to deviate. In this dissertation, I will show all three of these hold true for the CEO, and that there is significant empirical evidence indicating that this does indeed occur.

The strongest evidence for executives being able to influence firm capital structure decisions comes from survey data. Graham et al. (2013) report from a survey of executives that CEOs and CFOs claim that one of the two areas they have the most influence are capital structure, with the other area being mergers and acquisitions. Frank and Goyal (2007) provide more direct evidence by showing that differences among CEO fixed effects account for a large percentage of variation in corporate leverage. This dissertation provides further evidence by demonstrating that executive compensation partially determines the firm's leverage choice.

CEOs have a strong incentive to deviate from an optimal capital structure because

default is very costly for the current CEO. Very few CEOs survive the bankruptcy process and their outside option is normally significantly lower than their current wage. Betker (1995) reports that 91% of CEOs in office two years prior to a chapter 11 bankruptcy do not survive the bankruptcy process. Ayotte and Morrison (2009) find a similar percentage. Further, CEOs earning prospects are significantly decreased after being forced out. Eckbo and Thorburn (2003) find that in a sample of Swedish firms median income change for a CEO resulting from bankruptcy is -47%, and in a more recent working paper Eckbo et al. (2012) use a sample of U.S. firms undergoing bankruptcy to show that the median loss of future income due to undergoing bankruptcy is 2.7 times present income. Therefore, bankruptcy is very costly for the executives of the firm, substantially more costly than it is for the firm. This drives a wedge between the optimal capital structure from the perspective of the executive and the perspective of the firm.

It is a common assumption in the literature that CEOs possess incentives other than those of shareholder wealth maximization, such as empire building or perquisite consumption. The literature exploring optimal compensation schemes seeks to align the incentives of CEOs with shareholders, however, in most models, the firm is rarely able to achieve a first best solution through compensation contracts. Executive compensation typically consists of four main components, a fixed wage, various explicit performance incentives, and an equity stake in the company through both direct stock ownership and indirectly through options. In this dissertation, I will ignore the explicit performance incentive. I justify this by noting that explicit performance incentives tends to be minuscule compared to the magnitude of pay performance sensitivities due to stock and options in the data. However, the literature has ignored the presence of personal default costs for CEOs when examining the compensation contracts. Since executives are risk averse agents, this interaction is important.

A standard result in the compensation literature going back to Holmstrom and

Milgrom (1987) is that risk-sharing is fundamentally sub-optimal for a risk averse agent, and it's only necessary to provide correct incentives for effort. However, risk averse agents care not only about the amount of wealth, but also about what state of the world the payments arrive in. For an executive with a large component of compensation due to a pay for performance component of the contract, the state of the world under which he realizes his large negative wealth shock due to default is also the state of the world under which his pay for performance is very low. This means a heavy pay for performance based compensation contract further increases the perceived cost of firm default.

As the cost of default increases, if the CEO has some measure of control over firm leverage, the chosen leverage ratio decreases. This may lead to a lower than optimal leverage choice, though it's not a necessity. However, this interaction does lead to a chosen leverage ratio that is lower than what would otherwise be observed.

I build a model that demonstrates this intuition by showing that the risk adjusted probability of default increases as the probability of default increases. This leads to a negative relationship between PPS and firm leverage. Further, I show that the risk adjusted probability of default increases faster for more volatile firms, and therefore the negative relationship between PPS and firm leverage increases. The model also predicts an increasingly negative relationship as a firm becomes less likely to take disciplinary actions against executives for sub-optimal leverage choices. I then use firm level data on executive compensation to build a comprehensive estimate of the PPS for each CEO of approximately 1600 large publicly traded firms over the years 1992 through 2010, and I show that CEO PPS is significantly negatively related to firm leverage decisions, and that this effect is robust to alternative specifications of leverage and PPS. This result is graphically represented in figure 1.1, where a clear negative relationship between PPS and leverage is observed for market leverage. Figure 1.2 shows the same effect for market leverage. Figures B.1 and B.2 in Appendix

B show the deviations from average industry leverage as a function of PPS. Further, I am able to show that more volatile firms and firms that are less likely to punish the CEO for poor leverage decisions do indeed see a stronger negative relationship. CEOs of firms post-Sarbanes-Oxley are likely to face higher costs in default through an increased likelihood of criminal charges, and I demonstrate that the relationship between PPS and leverage increases after Sarbanes-Oxley is enacted. Further, I demonstrate that CEOs that have high PPS alter leverage adjustment speeds to maximize the amount of time spent in an underlevered state of the world, providing more causal evidence that CEOs are managing the debt levels of firms in response to pay performance sensitivity. Finally, I provide further evidence that CEO PPS is directly affecting firm leverage through using CEO tenure as an instrument for CEO equity holdings using a two stage least squares framework.

This effect is not only interesting for the implications it has for the assumption of shareholder wealth maximization being the goal of the firm, it is also economically significant. I estimate that approximately 1% of total firm value is destroyed through this channel, through forgoing the tax shield due to debt. However, this is not necessarily sub-optimal, ex-ante, for the firm. If the CEO has significant private information about the optimal firm leverage decision, then it may still be optimal for the CEO to make the leverage decision. Since PPS has been shown to increase firm performance over a variety of metrics, the overall effect of increased PPS on firm value may be positive. However, it's also possible that compensation committees are either not aware of or unable to affect this channel through which CEOs are extracting rents, and it may be optimal for the board of a firm to bring in an outside consultant to help set leverage targets in order to remove the CEO from this decision making process.

One might reasonably ask if the CEO will still have a significant exposure to firm value in the event of bankruptcy, if the CEO will have rid himself of his equity

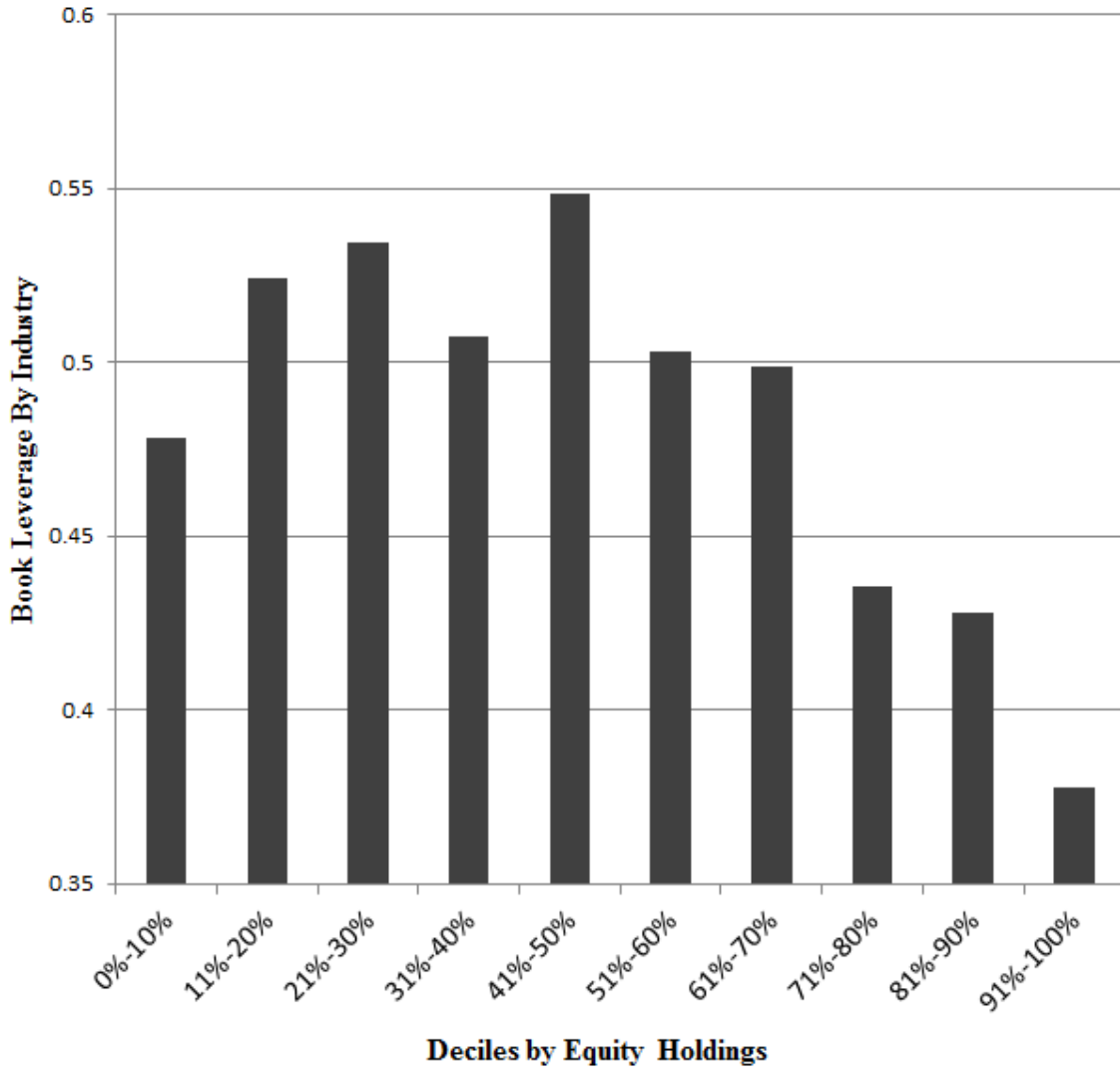


FIGURE 1.1: Book Leverage by Deciles of Pay Performance Sensitivity

position through either direct selling or using options to hedge. However, this is unlikely to be the case. It is very difficult for CEOs to sell stock when approaching a bankruptcy event due to restrictions on insider selling, and Eckbo et al. (2012) show that the median CEO equity value decreases by \$5 million over the course of a bankruptcy event. CEOs are also unlikely to be able to fully hedge their exposure to equity value. Option trading is subject to insider trading laws, and it is illegal to executives to short sell their own stock. However, Garvey (1997) argues that for sufficiently liquid option markets it is possible for managers to engage in purchasing

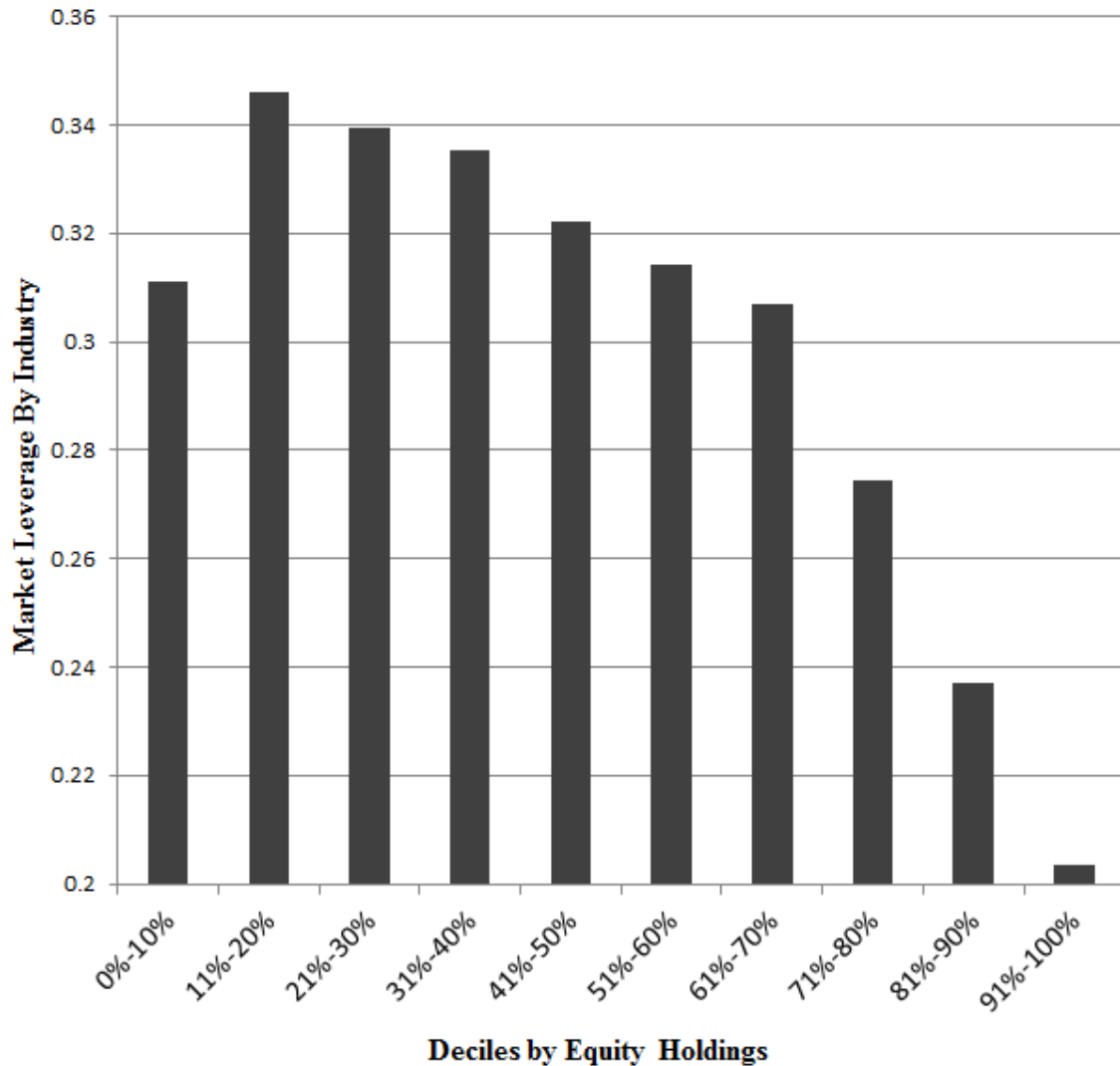


FIGURE 1.2: Market Leverage by Deciles of Pay Performance Sensitivity

put options in order to hedge. I test whether or not the liquidity of a firms option market decreases the magnitude of the relationship between PPS and firm leverage, and I find that it has a small, but statistically insignificant effect on the relationship.

The effect is non-trivial to document because of the relationship between pay performance sensitivity and firm leverage. My measure of pay performance sensitivity is closely related to the value of the CEOs equity holdings in the firm, and if I exclude options, it is identical to the value of the CEOs stock holdings in the firm. However, since firms only issue debt sporadically, there is a spurious relationship

between observed market leverage and the standard measures of PPS. Assuming that firms do possess a target leverage, the relationship between target market leverage and PPS should still hold, but it's not directly observable. You must, instead, use book leverage, which I do for this dissertation. For completeness, I document that the relationship is only stronger when one considers market leverage, however the regressions and univariate graphical analysis are partially spurious.

Further, there is question as to the appropriate measure for pay performance sensitivity. My story fundamentally relies on the idea that a significant portion of CEO wealth tied up in the equity of the firm is destroyed during a bankruptcy event, so it's important that PPS represents this wealth. Over the past two decades, options have become a significant component of a CEO's equity holdings in a firm. My full measure of PPS includes both the sensitivity of a CEO's stock holdings and option holdings to stock price movements. However, it's likely that the value of the option holdings, and therefore the sensitivity to stock price movements, will be destroyed long before the bankruptcy event as the stock price falls well below the strike price of many options in the CEO's portfolio. One should expect, then that the primary concern for the CEO is that of his stock holdings. In order to account for this criticism, I separate the value of option holdings from the stock holdings, and show that all results hold, and in most cases are strengthened, by considering only the stock holdings of the executive. I consider this evidence in support of the bankruptcy event being the primary consideration.

If the correlation between CEO personal wealth and default costs is driving this effect, the natural variable is the amount of CEO wealth in the firm. I use PPS as my main variable of interest throughout most of the dissertation, however, PPS is isomorphic to the value of stock holdings if options are excluded. However, I check all of my results using the value of stock and option holdings instead of PPS, and all results hold identically.

The empirically documented fact of this dissertation that the value of the CEOs stock holdings and not the value or sensitivity to stock price movements of her option holdings serves to distinguish between other competing explanations. The agency cost of debt theory argues that debt prices respond to CEO incentives, causing the relative price of debt and equity to change and potentially leading to leverage changes as a result (see Jensen and Meckling (1976) and John and John (1993)). This alternative explanation would argue that as CEOs are compensated with a larger portion of options, they have an incentive to increase the riskiness of a firm. Lenders, in equilibrium, understand that they will bear a disproportionate percentage of this risk due to this incentive to increase riskiness and raise relative borrowing costs for those CEOs. Since borrowing costs increase, the firm responds by using more equity financing and less debt financing. However, this explanation relies on option values being positively related to firm volatility. Since I find that stock holdings are the main driver of this effect, and stock value is not positively related to volatility, my results are inconsistent with this being the sole explanation. However, I do find a significant, though much smaller in magnitude, negative relationship between option holdings and firm leverage, this explanation may account for a part of the relationship between a measure of total PPS and firm leverage.

This dissertation is related to several different literatures. It's related to the literature on optimal executive compensation when there are agency problems in the vein of Holmstrom and Milgrom (1987). The literature that shows that compensation contracts have a direct effect on significant firm decisions such as Aggarwal and Samwick (2006), Aggarwal and Samwick (1999), and Denis et al. (1997) is closely tied to this research since I examine the effect of compensation on firm leverage. It's also related to a literature on CEO personal default costs such as Ayotte and Morrison (2009), Eckbo and Thorburn (2003), and Berk et al. (2010), since the alternative perspective of my results provide indirect evidence for costly personal

default for CEOs. The work this is most closely related to is Frank and Goyal (2007). They have a single table showing a negative relationship between PPS and leverage, but as discussed in the results section, the regressions in the specific table that reports this result is misspecified, and they don't discuss the table in any way. However, they show a significant CEO specific effect on firm leverage, but they look at a CEOs fixed effects as she moves firms, not the effect of compensation on firm leverage choice.

The dissertation is structured as follows. I first present my model and the solution. I then discuss the sources and construction of my data set. Next I report empirical results and provide some interpretation. I then conclude and indicate future directions for this line of research.

2

Model

The model is a principal agent model in the mold of Holmstrom and Milgrom (1987). Firm output is a stochastic function of both effort and leverage decision. The agent has the standard disutility of effort, but the agent also experiences a negative shock if the firm defaults. The principal provides a compensation contract that the agent then uses to make his optimal choice of effort and leverage.

In this model, I don't derive the optimal compensation contract. I think there is merit in deriving the optimal contract when the agent makes a leverage decision, but since I am primarily concerned with the agents response to a given compensation contract, I will always assume that the compensation contract is given. Depending on whether the compensation committees consider the affect on the firm leverage choice when they design the compensation contract, this may be the correct way to model this. This may happen if either the effect on firm value of a sub-optimal leverage decision is a second order effect, firms aren't aware of this effect, or the cost of providing incentives for the optimal leverage choice is too high, the last of which I find most likely. However, since this model is highly stylized, it can not speak directly to the magnitude of the effect. Whether compensation committees choose

the compensation contract optimally is irrelevant to the results of this dissertation. It does speak to the optimality of observed leverage ratios, but not to the relationship between PPS and leverage.

The principal can contract on the output of the firm, but following the standard agency problem, the effort level is unobservable. Leverage is observable, and the principal can contract directly on the leverage choice. However, I assume that the agent possesses a technology that makes him better suited to set leverage, so the executive only provides a linear incentive to increase leverage. In my model, the agent always strictly wants to decrease leverage relative to the principal, so the linearity has no significant restriction on the contract space. Explicit contracting on leverage is not observed in executive compensation contracts, but executives that make sub-optimal decisions for the firm potentially face disciplinary actions. I interpret the component of compensation dependent directly on leverage as a function of the probability of disciplinary actions by the board.

I don't model explicitly the agent specific technology for determining proper leverage choices because the model is highly stylized, but this can be justified by considering that the agent may be best positioned to predict the marginal tax rate that the firm is likely to experience, or he may understand better the firm specific cost of default. If there is an asymmetry in information related to the marginal tax rate or default costs, the optimal action for the principal to take would be to offset the agents tendency to underlever due to agent specific costs of default and allow the agent to directly set the leverage ratio. Again, this isn't explicit in my model since it distracts from the primary goal of providing intuition on the interaction between PPS and default costs.

I will make the assumption that the compensation contract is linear in firm performance. While this is not without loss of generality, observed compensation contracts are approximately linear, and this assumption is consistent with other work in this

literature (Holmstrom and Milgrom (1987), Holmstrom and Milgrom (1991), and Jin (2002)).

The major deviation from the standard model is the negative shock experienced by the agent in the case of default. This effect captures the observation that firm default is very costly for a CEO, much more so than for shareholders. I further assume that the principal faces no cost of default, other than the obvious loss of value. The equity holder of the firm must make up the shortfall to cover the cost of debt, but the principal will be risk neutral, so this will have no affect on the optimal ex-ante decision of the principal.

2.1 Details of Model

The agent is assumed to be risk averse with CARA utility, and it is assumed that the agent cannot diversify away firm specific risk. If the agent is allowed to diversify away firm specific risk, any contract that depends on firm performance will be immediately diversified away, and it will have no affect on the incentives of the agent. The model is single period, and since the agent has CARA utility I assume without loss of generality that the initial wealth of the agent is zero. The principal is risk neutral.

The firm value at the end of the period is given by $\pi = (1 + tL)x + \epsilon$ where x is the effort that the agent puts into the firm, L is the leverage ratio chosen by the firm, t is the marginal tax rate, and $\epsilon \sim N(0, \sigma^2)$ is a stochastic shock to the value of the firm. The only role of leverage in this model is to provide a tax shield on profits. Note that the effort choice x will be completely determined in equilibrium, there will be no information asymmetry, so leverage is well defined as a proportion of effort. However, if the realization of the firm's value cannot cover the amount of the amount of the debt, the firm enters default.

Default is the state in which $\pi - Lx < 0$, i.e. the value of the firm is less than the value of debt. This occurs for a sufficiently negative shock, $\epsilon < ((1 - t)L - 1)x$.

Note that even with no debt, $L = 0$, the firm still defaults if the value of the firm is less than zero. If the firm defaults, there is a negative payment of size d to the agent. The interpretation of this shock to agent wealth is the loss in lifetime income due to lowered future employment prospects. The assumption of default breaks the linearity of the model, and this adds considerable complexity to the solution.

The agents compensation contract is linear in the outcome of firm value and is given by $w = w_0 + \alpha\pi + \phi L$. The agent receives a fixed component w_0 , a percentage of profits α , and a payment to provide incentive for a higher leverage ratio at the rate ϕ . Though this contract is written as a positive payment for a larger choice of the leverage ratio, it is isomorphic to a negative payment for a low leverage ratio. As noted previously, we don't observe explicit clauses in compensation contracts for leverage, however, an executive that chooses a sub-optimal leverage ratio faces disciplinary actions with a probability that is proportional to the extent of the deviation. In my model, the tendency to deviate will always be in the negative direction, so the principal will only ever want to provide incentives for the agent to increase the leverage ratio. The model can be extended to punish both very high and very low leverage, but for simplicity I now assume that the principal only provides positive linear incentives for the managers leverage ratio.

The agent has CARA utility with a risk aversion of γ . The agent has a disutility of effort equal to $\frac{kx^2}{2}$. Further, if the firm defaults, the agent receives a negative shock to wealth of d . The total utility function of the agent becomes

$$U_A = e^{\gamma(w_0 + \alpha\pi + \phi L - \frac{kx^2}{2} - dI_{\text{default}})} \quad (2.1)$$

where I_{default} is a dummy variable which is one if the firm enters default. The agent has an outside option with value U_0 . Then the agents problem can be rewritten as

the following optimization:

$$\max_{x,L} \gamma(w_0 + \alpha(1 + tL)x + \phi L - \frac{kx^2}{2} - \frac{1}{2}\gamma\alpha^2\sigma^2) - \ln(e^{d\gamma}\Phi(\epsilon'_D) + (1 - \Phi(\epsilon'_D))) \quad (2.2)$$

s.t.

$$\gamma(w_0 + \alpha(1 + tL)x + \phi L - \frac{kx^2}{2} - \frac{1}{2}\gamma\alpha^2\sigma^2) - \ln(e^{d\gamma}\Phi(\epsilon'_D) + (1 - \Phi(\epsilon'_D))) \geq u_0 \quad (2.3)$$

where

$$\epsilon'_D = \frac{((1 - t)L - 1)x + \gamma\alpha\sigma^2}{\sigma} \quad (2.4)$$

The first term in equation (2.2) is exactly equivalent to a standard principal agent problem. However, the second term is unique to my model. Φ is the cumulative distribution function for the standard normal, and ϵ'_D looks almost like the standardized value of ϵ in default, however, it's not quite. Instead, it is the risk adjusted probability of default for the risk averse agent. Note that if $d = 0$, the second term becomes zero and this reduces exactly to a standard contracting problem. However, as it is, the problem is quite nonlinear, and difficult to arrive at closed solutions.

The leverage ratio L enters in the objective function in a few different places. First, there is a positive effect on profits through the tax benefit of debt that directly effects agents utility through the profit sharing portion of the agents compensation contract. Second, there is a direct positive effect through the incentive in the compensation contract for higher leverage. Third and finally, the leverage choice directly effects the probability of default, $\Phi(\epsilon'_D)$. This last effect causes the agent to reduce his leverage choice. His incentives are aligned with the principal to the extent that there is profit sharing, but the additional cost of default that the agent bears that the principal doesn't causes a wedge between the optimal leverage choice from the perspective of the principal and the observed leverage choice by the agent.

2.2 Solution

The introduction of non-linear default costs specific to the agent creates additional complexity to the solution of the model because it breaks the fundamental linearity of the contract. However, it is possible to determine some theoretical results that will provide intuition for the empirical results. As stated previously, I am interested in how the agent responds to his compensation contract when the principal isn't explicitly contracting on at least one choice variable. Further, I must assume that the agent is limited in his ability to remove the possibility of bankruptcy through effort alone. I do this through allowing his cost of effort, k , to be large. If the cost of effort is sufficiently large, he is unwilling to exert the effort necessary to avoid bankruptcy in all states of the world. Further, I am interested in the state of the world in which default is very costly, so I will only prove that my results hold for situations in which default is a very costly event. However, there is no strict lower bound on the necessary size of the default cost, so I will restrict my attention to arbitrarily large values.

Assumption 1. *Parameter values are subject to the following restrictions:*

1. *The personal cost of default for the agent, δ , is large.*
2. *The cost of effort of the agent, k , is large. Specifically, $k \geq \frac{1}{(1-t)\gamma\sigma^2}$.*
3. *Leverage, L , is in the set $[0, \infty]$, and parameters are such that the agent finds it optimal to choose $L > 0$.*
4. *The agent is risk averse, i.e. $\gamma > 0$.*
5. *The tax rate, t , is in the set $[0, 1]$.*

Given assumption 1, we can show the following theorems.

Theorem 2. *If the principal cannot contract on leverage, i.e. $\phi = 0$, leverage is negatively related to PPS, i.e. $\frac{dL}{d\alpha} < 0$.*

Proof. See Appendix A. □

Theorem 3. *If the principal cannot contract on leverage, i.e. $\phi = 0$, the relationship between leverage and PPS becomes more negative as stock volatility increases, $\frac{d^2L}{d\alpha d\sigma} < 0$.*

Proof. See Appendix A. □

Theorem 4. *The relationship between leverage and PPS becomes less negative if the principal does contract on leverage, $\frac{d^2L}{d\alpha d\phi} > 0$.*

Proof. See Appendix A. □

The above results provides insights into what we should see empirically, and it allows me to provide predictions inconsistent with other stories that might explain the observed negative relationship. First, theorem 2 implies that this is a causal relationship. While compensation contracts and firm characteristics relevant to the leverage decision are both determined contemporaneously and endogenously, my model predicts that large equity stakes in the firm should generate a negative relationship, not just be correlated with it. This is in contradiction to the explanation that high leverage has a disciplining effect on CEOs, and therefore is a substitute for high performance based compensation.

Theorem 3 provides an additional prediction inconsistent with non-causal relationship. A non-causal relationship would imply that if the variance of returns is controlled for, there should be no relationship between the magnitude of this effect and variance, since variance effects the chosen leverage, and the chosen leverage affects the compensation contract for the CEO, but there should be no additional correlation between leverage and performance based compensation.

Finally, theorem 4 provides a prediction that is distinct from the agency cost of debt. The agency cost of debt assumes that the executive ultimately affects the volatility of the firm through the choice of projects, but does not directly control leverage. The negative relationship is due to an increased cost of

Description of Data and Variables

My data comes from several different sources. My primary data set is the ExecuComp database of executive compensation measured at an annual frequency for each firm in the S&P 500, S&P mid-cap 400, and the S&P small cap 600 from 1992 to 2010. The dataset includes the compensation contract for the five most highly paid executives in the firm, including salary, bonus, stock grants, and option grants. I am able to construct from this database an accurate measure of CEO PPS.

The pay performance sensitivity is not a trivial thing to measure because firms do not report the strike price and time to maturity of options not granted during the current fiscal year, though they do report all information for options granted during the current fiscal year. To calculate pay performance sensitivity one must either look at all previous option grants to the executive or estimate the value of those previous grants. To examine all previous option grants, several years of data are needed, and since ExecuComp is a fairly recent database, this would be prohibitively expensive in terms of discarded data. The method I use is the estimation procedure in Core and Guay (2002). In this procedure, the existing grants are assumed to have a certain strike and time to maturity based on the most recent option grants. This allows me

to calculate the total pay performance sensitivity for the executive with only a single year of data. Core and Guay (2002) are able to show that this procedure captures more than 99% of the the variation in option portfolio value and sensitivities. I also look at the pay performance sensitivity due to the current stock holdings of the executive without consideration of the option portfolios, and this value can be acquired directly and accurately as it is equivalent to the value of the stock holdings for the executive. My variable of interest is the natural logarithm of pay performance sensitivity (following Brockman et. al. (2010)), since pay performance sensitivity is very right skewed.

As a robustness test, I calculate the value of the option and stock grants directly and I use the logarithm of those values in place of the pay performance sensitivity discussed above. Since, pay performance sensitivity due to stock grants is a linear transformation of the value of those stock grants, results are identical for the stock grants. However, while the pay performance sensitivity due to the option grants is closely related to the value of the option grants, it is not identical. When I replace pay performance sensitivity by the value for option grants, though all results are qualitatively unchanged.

Though the literature uses the term pay performance sensitivity to describe the variable of interest, it is closer to the amount of wealth the CEO has invested in the firm.

I combine pay performance sensitivity with the Compustat annual database for firm level information to calculate firm level variables. I use market leverage, book leverage, total assets, industry market and book leverage, and the ratio of property plants and equipment to total assets (a measure of tangibility). See table (3.1) for the details of the calculation of the variables.

I also use several other measures of firm state. I use two measures of corporate governance to proxy for the probability of termination due to sub-optimal leverage

Table 3.1: Variable Definitions

Variable (Data Source)	Variable Definition
Compustat	
Total Assets	Assets - Total (AT)
Book Equity	Stockholders Equity - Total (SEQ) + Deferred Taxes and Investment Tax Credit (TXDITC) - Preferred/Preference Stock (Capital) - Total(PSTK) if (PSTK) missing then Preferred Stock Redemption Value (PSTKRV) if (PSTKRV) missing then Preferred Stock Liquidating Value (PSTKL)
Book Debt	Assets - Total (AT) - Book Equity
Book Leverage	Book Debt / Assets - Total (AT)
Market Leverage	Book Debt/ (Common Shares Outstanding (CSHO) * Price Close - Annual Fiscal year (PRCC.F) + Book Debt)
Market-to-Book	(Common Shares Outstanding (CSHO) * Price Close - Annual Fiscal year (PRCC.F) / Book Debt (BD)) / Assets - Total (AT)
Abnormal Earnings	(Income Before Extraordinary Items Adjusted for Common Stock Equivalents in $t + 1$ (IBADJ) - Income Before Extraordinary Items Adjusted for Common Stock Equivalents (IBADJ) in t) / (Common Shares Outstanding (CSHO) * Price Close - Annual Fiscal year (PRCC.F))
ROA	Operating Income Before Depreciation (OIDBP) / Assets - Total (AT)
CRSP	
Unlevered Volatility	One year rolling window on monthly returns, unlevered
ExecuComp	
Salary and Bonus	Salary (SALARY) + Bonus (BONUS)
PPS	See Core and Guay 2002 for estimation procedure
PPS_stock	Price Close - Annual Fiscal Year (PRCC.F) * Shares Owned - Options Excluded (SHROWN_EXCL_OPTS) / 100
CEO Ownership	Shares Owned - Options Excluded (SHROWN_EXCL_OPTS) / Total Share Outstanding (SHRSOUT)
Measures of Entrenchment	
Gompers index	24 anti-takeover provisions index by Gompers, Ishii and Metrick (2003)
Bebchuk index	6 anti-takeover provisions index by Bebchuk, Cohen, and Farrell (2004)
FED	
Risk-free rate	Average T-bill rate

choice. I use the entrenchment index provided by Bebchuk et. al. 2009, which is a measure of the number of entrenchment provisions the CEO has in place. Similarly, I use the governance index provided by Gompers et. al. 2003, which is another measure of entrenchment provisions. Both of these measures have been shown to be correlated with value destroying actions by the CEO. I then generate a dummy variable for firms that are in the top 20% of firms in terms of take-over provisions. I also calculate the historical volatility using a one year rolling window of monthly stock returns. I then calculate an indicator variable for firms that are in the top or bottom 20% of volatility. I also measure the tenure of the CEO sitting in that year because Eckbo, Thorburn, and Wang (2012) show that CEO tenure is a predictor of the magnitude of the negative outcome in the case of default, with the intuition being that a CEO with longer tenure is more likely to be blamed for the bankruptcy event.

The final variable that I calculate is a measure of under or over-leverage. It is difficult to measure explicitly a firms deviation from optimal leverage, however, Binsbergen et. al. (2010) provides a method to calculate the marginal cost of an additional unit of debt relatively simply. Then I compare the marginal benefit due to the tax shield of debt using firms effective marginal tax rate calculated as in Graham and Mills (2008). I then take the deviation of the marginal cost of debt from the marginal benefit of debt and I assume that firms for which the marginal benefit of debt far exceeds the marginal cost of debt are underlevered, and vice versa if the marginal cost of debt far exceeds the marginal benefit. I then use this to determine the 20% most overlevered and 20% most underlevered firms. While this is a relatively crude method through which to calculate over or underleverage, it should be sufficient for the relatively coarse use of calculating the dummy variable.

I require that all observations must have data for market leverage, book leverage, and pay performance sensitivity. This leaves 12,611 firm year observations. However,

as a control I use a measure of tangibility, property plants and equipment scaled by total assets, which has only 12,544 overlapping observations, so for regressions with controls included there are 12,544 observations. My sample is positively skewed on size, leverage, and compensation metrics. I drop all regulated industries (two digit SIC code 49) and financial services industries (two digit SIC codes 60-69), following the literature. For the descriptive statistics of my sample see table (3.2).

Since components of compensation are positively skewed, I take the natural logarithm of pay performance sensitivity consistent with Brockman et. al. (2010). I then standardize all variables.

Table 3.2: Summary Statistics

VARIABLE	N	Mean	Median	S.D.	25 th %ile	75 th %ile
Investment and Financial Characteristics						
Market Leverage	12,333	.300	.267	.201	.137	.426
Book Leverage	12,333	.474	.469	.252	.313	.600
Total Assets (in billions)	12,333	5.33	.936	23.5	.370	3.00
Property Plants and Equipment (in billions)	12,333	3.03	.424	11.7	.134	1.68
Property Plants and Equipment/Assets	12,333	.552	.467	.376	.264	.764
Market to Book	12,333	2.200	1.68	1.751	1.27	2.49
Unlevered Volatility	12,333	.085	.069	.064	.046	.106
ROA	12,333	.143	.147	.289	.098	.202
Abnormal Earnings	12,333	-.002	.007	3.50	-.011	.024
CEO compensation						
Salary and Bonus (in millions)	12,333	1.24	.873	1.65	.530	1.46
PPS (in millions)	12,333	1.15	.233	6.45	.090	.647
PPS stock only (in millions)	12,333	.860	.064	6.20	.016	.257
CEO Ownership	12,333	.031	.004	.073	.001	.021
Measures of Entrenchment						
Bebchuk index	7,436	2.44	3	1.28	2	3

Empirical Methodology and Results

This section tests the predictions of the model. My regressions are somewhat opposite of the standard literature (Aggarwal and Samwick (1999), Jin (2002)) that uses pay performance sensitivity as the dependent variable. However, I am attempting to explain the observed leverage ratios as a function of CEO compensation, while both Aggarwal and Samwick (1999) and Jin (2002) are interested in explaining the determinants of executive compensation.

4.1 Regression Estimation

First, I estimate the following equations

$$\text{Leverage} = \alpha + \beta \ln(\text{PPS}) + \beta_c \text{Controls} + \epsilon \quad (4.1)$$

All regressions include both firm and year fixed effects. I estimate most regressions with and without controls. The controls that I include are the four reliable factors for predicting cross-sectional leverage as described by Frank and Goyal (2008), median industry leverage, log of assets, market-to-book, and a measure of tangibility as well as addition controls for return on assets, abnormal earnings, and percent of equity

Table 4.1: Regression from Equation (4.1) with Full PPS Measure

VARIABLES	(1) Market Leverage	(2) Market Leverage	(3) Book Leverage	(4) Book Leverage
ln(PPS)	-0.278*** [-9.733]	-0.257*** [-8.041]	-0.174*** [-4.906]	-0.160*** [-3.841]
ln(Assets)		0.311*** [6.864]		0.0470 [0.987]
Market to Book		-0.165*** [-5.520]		0.121 [1.086]
PP&E/Assets		0.0732** [2.144]		0.160*** [3.898]
Mean Industry Market Leverage		0.269*** [14.32]		
Mean Industry Book Leverage				0.154*** [3.594]
Abnormal Earnings		-0.00519 [-0.828]		-0.0156** [-2.391]
CEO Ownership		0.139*** [5.792]		0.0796*** [3.725]
ROA		-0.0686* [-1.894]		-0.269*** [-5.472]
Constant	-0.264*** [-8.990]	-0.115*** [-3.711]	-0.197*** [-6.617]	-0.191*** [-4.944]
Year and Firm Fixed Effects	Yes	Yes	Yes	Yes
Observations	12,333	12,333	12,333	12,333
Adj.R-squared	0.754	0.806	0.596	0.694

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

held by the CEO. All regressions are robust and standard errors are clustered at the firm level.

Tables (4.1), (4.2), and (4.4) reports results for the regression in equation (4.1). In table (4.1), I report results for the regression for the pay performance sensitivity including both stock and options in the calculation. The important thing to note in this table is the large and statistically significant regression coefficient for PPS. The regression coefficient is larger and more significant than all of the controls. Note that all variables are standardized, so the interpretation is one standard deviation in the log of PPS results in a .160 standard deviation decrease in the book leverage.

One thing that is important to notice in table (4.1) is the large difference between the effect on market leverage and book leverage. This is due to two factors, 1) book leverage in general is less predictable than market leverage (Frank and Goyal

Table 4.2: Regression from Equation (4.1) with PPS Measure in Components

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage
ln(PPS_stock)	-0.101*** [-4.211]	-0.0992*** [-4.176]	
ln(PPS_option)		-0.0495* [-1.955]	-0.0546** [-2.124]
ln(Assets)	-0.00972 [-0.205]	0.00277 [0.0577]	-0.0367 [-0.713]
Market to Book	0.101 [0.925]	0.104 [0.951]	0.0871 [0.817]
PP&E/Assets	0.168*** [4.133]	0.167*** [4.071]	0.178*** [4.284]
Mean Industry Leverage Book	0.158*** [3.566]	0.156*** [3.568]	0.161*** [3.653]
Abnormal Earnings	-0.0205*** [-2.824]	-0.0201*** [-2.812]	-0.0220*** [-2.838]
CEO Ownership	0.0560*** [3.147]	0.0477*** [2.637]	-0.00159 [-0.0898]
ROA	-0.274*** [-5.318]	-0.274*** [-5.355]	-0.276*** [-5.296]
Constant	-0.169*** [-4.637]	-0.182*** [-4.864]	-0.193*** [-4.954]
Year and Firm Fixed Effects	Yes	Yes	Yes
Observations	12,333	12,333	12,333
Adj.R-squared	0.692	0.693	0.689

Robust t-statistics in brackets
 *** p<0.01, ** p<0.05, * p<0.10

(2008)) and 2) the regression on market leverage has a spurious element in it. The calculation of PPS includes both stock and options, but the PPS of options, the delta, is a function of stock price. As stock price increases, market leverage decreases and options become more in the money, increasing the delta of the options. This causes a mechanical negative correlation between PPS and market leverage. For this reason, outside of tables (4.1), (4.2), and (4.4), I will do all further analysis using only book leverage, however all results both hold and are strengthened with market leverage as the dependent variable.

While table (4.1) only uses log of PPS, table (4.2) breaks the PPS into the component due to stock holding by the CEO and that due to the option holdings by the CEO. Note that for both market leverage and book leverage, the coefficients for each component of PPS are both significant and negative. However, the magnitude

Table 4.3: Regression from Equation (4.1) with Black-Scholes Value for Options

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage
ln(PPS_stock)	-0.101*** [-4.219]	-0.0808*** [-4.179]	
ln(BS Option Value)		-0.0186** [-2.369]	-0.0191** [-2.426]
ln(Assets)	-0.0105 [-0.222]	-0.0207 [-0.426]	-0.0497 [-0.964]
Market to Book	0.100 [0.919]	0.0958 [0.887]	0.0839 [0.789]
PP&E/Assets	0.167*** [4.116]	0.171*** [4.157]	0.180*** [4.330]
Mean Industry Book Leverage	0.157*** [3.569]	0.158*** [3.595]	0.162*** [3.658]
Abnormal Earnings	-0.0205*** [-2.823]	-0.0201*** [-2.734]	-0.0223*** [-2.822]
CEO Ownership	0.0565*** [3.186]	0.0415** [2.496]	0.00496 [0.302]
ROA	-0.274*** [-5.322]	-0.275*** [-5.302]	-0.276*** [-5.265]
Constant	-0.170*** [-4.645]	-0.174*** [-4.725]	-0.182*** [-4.789]
Observations	12,333	12,333	12,333
Adj.R-squared	0.692	0.692	0.689

Robust t-statistics in brackets
 *** p<0.01, ** p<0.05, * p<0.10

of the coefficient for the PPS due to stock is approximately twice that of the PPS due to options. Further, the coefficient for PPS due to stock has a much larger t-statistic, especially when both are estimated simultaneously. PPS due to stock seems to be the main contributor to this effect.

One criticism could be that PPS due to options isn't capturing the wealth at risk of the CEO well, since PPS measures sensitivity of value to stock price movements. In order to test for this, I replace PPS due to options with the Black-Scholes value of option holdings for the CEO. Note that since PPS due to stock is isomorphic to the stock holdings, I still use PPS due to stock as the relevant variable. Results are reported in table 4.3. I see consistent results with table 4.2, however the difference in magnitude between the relationship between leverage and stock holdings and the relationship between leverage and option holdings is significantly increased, with the coefficient on option holdings being less than 25% of the coefficient on stock

Table 4.4: Regression from Equation (4.1) with PPS Measure from Stock Only

VARIABLES	(1) Market Leverage	(2) Market Leverage	(3) Book Leverage	(4) Book Leverage
ln(PPS_stock)	-0.175*** [-11.21]	-0.148*** [-9.427]	-0.112*** [-5.105]	-0.101*** [-4.238]
ln(Assets)		0.204*** [4.872]		-0.0163 [-0.338]
Market to Book		-0.199*** [-6.173]		0.100 [0.924]
PP&E/Assets		0.0847** [2.490]		0.167*** [4.102]
Mean Industry Leverage Market		0.292*** [15.50]		
Mean Industry Leverage Book				0.159*** [3.634]
Abnormal Earnings		-0.0129* [-1.764]		-0.0203*** [-2.810]
CEO Ownership		0.0958*** [4.947]		0.0571*** [3.212]
ROA		-0.0764** [-1.972]		-0.273*** [-5.326]
Constant	-0.191*** [-6.628]	-0.0860*** [-2.715]	-0.152*** [-5.551]	-0.172*** [-4.680]
Year and Firm Fixed Effects	Yes	Yes	Yes	Yes
Observations	12,333	12,333	12,333	12,333
Adj.R-squared	0.737	0.795	0.589	0.690

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

holdings. I have re-estimated all tests using the Black-Scholes option value (though not reported here) and results are identical to those reported.

Finally, table (4.4) reports results only looking at the pay performance sensitivity due to the stock holdings of the CEO. I report these results to handle the criticism that default isn't the main concern for executives with large option portfolios since once options are sufficiently far out of the money, the value is practically zero. By this logic, the important component of PPS is that due to the stock held by the executive. However, for options deeply in the money, or options that may be exercised, there is likely to be an affect on the leverage choice. For this reason, for all future regressions, I include both the specification with the total PPS measure and the PPS due solely to the stock holdings of the CEO. Further, I don't include these results in this draft, but if I regress market leverage only on pay performance sensitivity due strictly to

stock holdings by the executive, the regression is no longer spurious since the spurious component is due to option valuations. All of my results are strictly stronger if I replace book leverage for market leverage and regress on PPS due to stocks only.

One of the contributions of this dissertation is that I document clearly and with a full specification of controls the effect of PPS on leverage. This effect was reported in a single table in Frank and Goyal (2007), but they use market leverage and a PPS measure that includes stock and options which has the issues of spuriousness. In addition, they don't use the logarithm of PPS, so their independent variable is very right skewed and not consistent with the literature. Finally, they don't look at the robustness of the result, nor even discuss their finding. In tables (4.1), (4.2), and (4.4), I am able to show that this effect is significant both statistically and in magnitude and that the effect is robust.

This result is counter intuitive given the current assumption prevalent in the literature that higher PPS results in a better alignment of managerial actions with shareholder preferences and that firms are underlevered. Widespread underleverage is unnecessary to my results, and there has been recent work that has questioned the result that firms are underlevered, but if the average firm is underlevered, the CEO is acting opposite of the interests of shareholders as PPS increases. Nor is the magnitude minor. For a one standard deviation increase in the log of PPS, the book leverage for an average firm decreases by .04, or 15% of the standard deviation of firm leverage.

While the results in tables (4.1), (4.2), and (4.4) are documentation of the negative relationship between PPS and leverage, they say very little about the underlying cause of the relationship. To further test the predictions of my model, I run the following regression

$$\text{Leverage} = \alpha + \beta_1 \ln(\text{PPS}) + \beta_2 \text{Dummy} \ln(\text{PPS}) + \beta_3 \text{Dummy} + \beta_c \text{Controls} + \epsilon \quad (4.2)$$

This is a diff-in-diff specification, but instead of looking at a dummy that will indicate an event, I examine a dummy that indicates a specific characteristic of the firm. I will examine firms with high and low volatility, high or low relative leverage, and good and bad corporate governance. If the interpretation of a standard regression coefficient is that of a derivative at a point, the interpretation of the regression coefficient β_2 should be that of a second derivative. β_2 gives the difference of the coefficient for the sample as a whole from the subsample indicated by the dummy variable. If β_2 is statistically significant then the response of the leverage choice to PPS changes as the variable that is used in the dummy changes. For all regressions, I will look at the top 20% and bottom 20% for each variable of interest.

Tables (4.5) and (4.6) look at the firms with the 20% highest and lowest unlevered volatility. Since, the volatility is unlevered, this is a proxy for how likely the firm is to hit bankruptcy. If the CEO is primarily concerned with the personal costs imposed on him during bankruptcy and, specifically, how they interact with his total wealth, as a firm becomes more volatile the CEO should want to decrease firm leverage even more, as described in the model. This is exactly what is observed in both tables (4.5) and (4.6). For firms with very high volatility the sensitivity decreases (becomes more negative) significantly. Similarly, for low volatility firms, those that likely do not have a high probability of bankruptcy, the sensitivity becomes significantly less negative. The results are robust to looking at PPS from only stocks, as table (4.6) shows.

Another prediction of my model is that CEOs with a greater incentive to choose proper leverage will reduce leverage less in response to an increase in PPS, i.e. if ϕ in my model is higher, than the leverage chosen for a given PPS will be higher. To test this, I examine firms that are objectively either under levered or over levered. As described in the preceding section, I calculate the difference between the marginal cost and the marginal benefit of an additional dollar of debt. I then say that firms

Table 4.5: Regression from Equation (4.2) with Full PPS Measure and Volatility as the Dummy

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage	(4) Book Leverage
ln(PPS)	-0.129*** [-4.959]	-0.137*** [-3.438]	-0.183*** [-4.855]	-0.166*** [-3.854]
High Volatility Dummy*ln(PPS)	-0.161** [-2.349]	-0.0906*** [-3.176]		
High Volatility Dummy	0.0560** [2.157]	0.0331 [1.397]		
Low Volatility Dummy*ln(PPS)			0.0606*** [2.605]	0.0385** [2.020]
Low Volatility Dummy			-0.0780*** [-3.577]	-0.0713*** [-3.274]
ln(Assets)		0.0588 [1.242]		0.0575 [1.213]
Market to Book		0.126 [1.126]		0.123 [1.099]
PP&E/Assets		0.157*** [3.864]		0.162*** [3.934]
Mean Industry Book Leverage		0.152*** [3.595]		0.153*** [3.571]
Abnormal Earnings		-0.0124** [-2.121]		-0.0155** [-2.366]
CEO Ownership		0.0765*** [3.586]		0.0787*** [3.688]
ROA		-0.265*** [-5.458]		-0.269*** [-5.482]
Constant	-0.196*** [-6.887]	-0.185*** [-4.916]	-0.169*** [-5.661]	-0.162*** [-4.259]
Year and Firm Fixed Effects	Yes	Yes	Yes	Yes
Observations	12,333	12,333	12,333	12,333
Adj.R-squared	0.600	0.696	0.597	0.695

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

whose marginal benefit greatly exceeds the marginal cost are under levered and vice versa. If a firm is greatly under levered it is likely that the board will take disciplinary actions against the CEO, since the sub-optimal choice should be evident, so the ϕ in my model should be higher. Tables (4.7) and (4.8) look at this by examining the 20% most under and over levered firms. Note that for firms that are under levered, the coefficient is large and statistically significant. For firms that are over levered, the coefficient is generally negative, in line with the predictions of the theory, but not statistically significant. The effect is robust, and even becomes stronger, when looking solely at PPS due to stock holdings only.

Table 4.6: Regression from Equation (4.2) with PPS Measure from Stock Only and Volatility as the Dummy

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage	(4) Book Leverage
ln(PPS_stock)	-0.0848*** [-4.989]	-0.0842*** [-3.786]	-0.117*** [-5.011]	-0.104*** [-4.182]
High Volatility Dummy*ln(PPS_stock)	-0.0960** [-2.293]	-0.0603*** [-2.594]		
High Volatility Dummy	0.0765*** [3.043]	0.0438* [1.861]		
Low Volatility Dummy*ln(PPS_stock)			0.0418* [1.858]	0.0219 [1.068]
Low Volatility Dummy			-0.0820*** [-3.617]	-0.0711*** [-3.148]
ln(Assets)		-0.0104 [-0.217]		-0.00756 [-0.158]
Market to Book		0.102 [0.940]		0.101 [0.933]
PP&E/Assets		0.165*** [4.102]		0.169*** [4.140]
Mean Industry Book Leverage		0.158*** [3.650]		0.158*** [3.618]
Abnormal Earnings		-0.0189*** [-2.734]		-0.0204*** [-2.803]
CEO Ownership		0.0552*** [3.120]		0.0563*** [3.159]
ROA		-0.272*** [-5.317]		-0.273*** [-5.336]
Constant	-0.163*** [-5.955]	-0.176*** [-4.837]	-0.127*** [-4.529]	-0.147*** [-3.984]
Year and Firm Fixed Effects	Yes	Yes	Yes	Yes
Observations	12,333	12,333	12,333	12,333
Adj.R-squared	0.591	0.691	0.590	0.691

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

The next table, (4.9), looks at how the sensitivity of leverage to CEO PPS changes as corporate governance changes. The Bechuk, Cohen, and Farrell index measures how many out of a set of 6 anti-takeover provisions the firm has in place. As the number of entrenchment indices increase, the CEO has less likelihood of being forced out short of bankruptcy, which typically ends in termination without the takeover provisions being activated. The Bebchuk dummy indicates the firms in the top 20% of anti-takeover provisions. As the number of anti-takeover provisions increase, the likelihood of the CEO facing termination for a sub-optimal leverage choice decreases, i.e. ϕ in the model decreases. Therefore, the sensitivity of leverage to PPS should

Table 4.7: Regression from Equation (4.2) with Full PPS Measure and Relative Over or Under Leverage as the Dummy

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage	(4) Book Leverage
ln(PPS)	-0.167*** [-2.877]	-0.130** [-2.422]	-0.123** [-2.490]	-0.103** [-2.269]
Most Underlevered*ln(PPS)	0.110** [2.207]	0.0375 [1.140]		
Most Underlevered	-0.232*** [-6.767]	-0.219*** [-6.274]		
Least Underlevered*ln(PPS)			-0.0781 [-1.152]	-0.0777 [-1.320]
Least Underlevered			0.258*** [5.019]	0.234*** [4.800]
ln(Assets)		0.0340 [0.435]		0.0194 [0.250]
Market to Book		0.151 [1.097]		0.146 [1.078]
PP&E/Assets		0.0706 [1.198]		0.0752 [1.276]
Mean Industry Book Leverage		0.194*** [4.363]		0.185*** [4.150]
Abnormal Earnings		0.238 [0.815]		0.254 [0.885]
CEO Ownership		0.0743*** [2.582]		0.0690** [2.428]
ROA		-0.229*** [-4.685]		-0.231*** [-4.813]
Constant	-0.0712* [-1.687]	-0.0663 [-1.428]	-0.159*** [-3.990]	-0.159*** [-3.325]
Year and Firm Fixed Effects	Yes	Yes	Yes	Yes
Observations	7,421	7,421	7,421	7,421
Adj.R-squared	0.651	0.766	0.653	0.769

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

become more negative. Tables (4.9) show that this is generally the case, however the coefficients are not significant. Note that the sample size is much smaller for Tables (4.9) due to the need to have information on the entrenchment indices, and both indices are fairly crude measures of managerial entrenchment, so it's not surprising that there is no statistical significance.

Table (4.10) attempts to control for the magnitude of the negative outcome in default by looking at CEO tenure. Eckbo, Thorburn, and Wang (2012) show that more senior CEOs are more likely to experience a large negative shock to lifetime earning. The explanation that they give is that more senior CEOs are more likely

Table 4.8: Regression from Equation (4.2) with PPS Measure from Stock Only and Relative Over or Under Leverage as the Dummy

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage	(4) Book Leverage
ln(PPS_stock)	-0.0800** [-2.417]	-0.0677** [-2.464]	-0.0458 [-1.541]	-0.0454** [-2.040]
Most Underlevered*ln(PPS_stock)	0.0689** [2.210]	0.0240 [1.011]		
Most Underlevered	-0.239*** [-7.114]	-0.222*** [-6.289]		
Least Underlevered*ln(PPS_stock)			-0.107* [-1.791]	-0.0958* [-1.800]
Least Underlevered			0.274*** [5.086]	0.247*** [4.805]
ln(Assets)		-0.0108 [-0.130]		-0.0209 [-0.256]
Market to Book		0.134 [1.000]		0.130 [0.985]
PP&E/Assets		0.0754 [1.282]		0.0809 [1.381]
Mean Industry Book Leverage		0.203*** [4.634]		0.194*** [4.444]
Abnormal Earnings		0.231 [0.773]		0.251 [0.861]
CEO Ownership		0.0531** [2.270]		0.0485** [2.047]
ROA		-0.231*** [-4.588]		-0.233*** [-4.687]
Constant	-0.0145 [-0.394]	-0.0394 [-0.898]	-0.115*** [-3.080]	-0.137*** [-3.084]
Year and Firm Fixed Effects	Yes	Yes	Yes	Yes
Observations	7,421	7,421	7,421	7,421
Adj.R-squared	0.645	0.763	0.649	0.766

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

to be blamed for the decisions that led the firm to bankruptcy, and therefore be blamed by shareholders and removed during the bankruptcy event or be tainted when entering the job market again. I create a dummy for the top 20% most senior CEOs in my sample by year and interact that with PPS. The interaction causes the effect to become consistently more negative for every specification, however the interaction term is not significant at the 10% level.

Table (4.11) examines whether CEOs that are likely to be able to hedge are significantly different than CEOs who can't hedge. Following Gao (2010), I consider a CEO that can hedge as one whose firm has high option volatility. My threshold is

Table 4.9: Regression from Equation (4.2) with Full and Stock Specific PPS Measure and Proxy for Governance as the Dummy

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage	(4) Book Leverage
ln(PPS)	-0.116*** [-4.268]	-0.101*** [-3.261]		
Bebchuk Dummy*ln(PPS)	-0.0831 [-1.607]	-0.0672 [-1.293]		
ln(PPS_stock)			-0.0644*** [-3.731]	-0.0573*** [-3.065]
Bebchuk Dummy*ln(PPS_stock)			-0.0676 [-1.298]	-0.0451 [-0.946]
Bebchuk Dummy	0.0441 [0.917]	0.0586 [1.228]	0.0419 [0.849]	0.0587 [1.209]
ln(Assets)		-0.000284 [-0.00400]		-0.0417 [-0.596]
Market to Book		-0.00961 [-0.336]		-0.0275 [-1.028]
PP&E/Assets		0.0422 [0.707]		0.0494 [0.823]
Mean Industry Book Leverage		0.197*** [7.183]		0.204*** [7.378]
Abnormal Earnings		0.185 [0.924]		0.168 [0.797]
CEO Ownership		0.0735*** [3.054]		0.0554** [2.580]
ROA		-0.195** [-2.520]		-0.217*** [-2.780]
Constant	-0.101*** [-2.724]	-0.108*** [-2.751]	-0.0642* [-1.766]	-0.0875** [-2.283]
Year and Firm Fixed Effects	Yes	Yes	Yes	Yes
Observations	7,436	7,436	7,436	7,436
Adj.R-squared	0.766	0.781	0.762	0.779

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

the top 20% of firms with option volatility, though the results are unchanged for other choices. I only report results for PPS from stock only since the option positions for the CEO are reported, there should be no effect on the PPS measure. Table (4.11) shows that CEOs that can hedge do have a smaller relationship between PPS and leverage, however it's statistically insignificant and a small fraction of the observed relationship. From this I conclude that CEOs do hedge, but not the entirety of their equity exposure. This is consistent with previous literature.

If default risk is driving this relationship, the ideal test would be to look at situations of heterogeneity in default risk. Unfortunately, there are no accepted

Table 4.10: Regression from Equation (4.2) with PPS Measure from Stock Only and Length of Tenure as the Dummy

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage	(4) Book Leverage
ln(PPS)	-0.195*** [-4.346]	-0.161*** [-3.663]		
Tenure Dummy*ln(PPS)	-0.00642 [-0.137]	-0.0325 [-0.807]		
ln(PPS_stock)			-0.122*** [-4.319]	-0.100*** [-3.732]
Tenure Dummy*ln(PPS_stock)			-0.0397 [-1.007]	-0.0511 [-1.500]
Tenure Dummy	-0.00668 [-0.0440]	-0.0646 [-0.358]	-0.145 [-1.065]	-0.154 [-0.846]
ln(Assets)		0.0153 [0.311]		-0.0471 [-0.934]
Market to Book		0.130 [1.117]		0.110 [0.970]
PP&E/Assets		0.196*** [5.654]		0.202*** [5.930]
Mean Industry Book Leverage		0.152*** [3.438]		0.157*** [3.476]
Abnormal Earnings		-0.0160** [-2.561]		-0.0208*** [-3.029]
CEO Ownership		0.0707*** [3.184]		0.0506** [2.550]
ROA		-0.267*** [-5.493]		-0.271*** [-5.358]
Constant	-0.208*** [-3.712]	-0.178*** [-4.129]	-0.126** [-2.555]	-0.137*** [-3.190]
Year and Firm Fixed Effects	Yes	Yes	Yes	Yes
Observations	11,454	11,454	11,454	11,454
Adj.R-squared	0.589	0.694	0.581	0.690

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

proxies for default risk. However, Brickley (2003) and Brickley (2006) show that top executives are significantly more likely to face criminal penalties post the enactment of the Sarbanes-Oxley Act of 2002 (hereafter referred to as SOX) due to bankruptcy. If an executive faces an increased risk of criminal litigation, the personal cost of default should increase. To test for this, I interact my PPS measure with a dummy for post-SOX. I see a significant increase in the magnitude of the effect post-SOX, consistent with default cost driving the relationship.

Since the effect is fundamentally driven by the probability of bankruptcy, we should expect to see a differential in the effect for firms of various sizes. This is due

Table 4.11: Regression from Equation (4.2) with PPS Measure from Stock Only and Option Volatility as the Dummy

VARIABLES	(1) Book Leverage	(2) Book Leverage
ln(PPS_stock)	-0.128*** [-3.385]	-0.104*** [-3.901]
Option Volume Dummy*ln(PPS_stock)	0.0307 [0.887]	0.00593 [0.250]
Option Volume Dummy	-0.169*** [-5.358]	-0.103*** [-3.607]
ln(Assets)		0.00681 [0.141]
Market to Book		0.103 [0.947]
PP&E/Assets		0.163*** [4.020]
Mean Industry Book Leverage		0.157*** [3.604]
Abnormal Earnings		-0.0201*** [-2.788]
CEO Ownership		0.0560*** [3.196]
ROA		-0.273*** [-5.346]
Constant	-0.151*** [-5.554]	-0.162*** [-4.437]
Year and Firm Fixed Effects	Yes	Yes
Observations	12,333	12,333
Adj.R-squared	0.591	0.691

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

to the empirical fact that large firms rarely fail, while small firms are much more likely to fail. In order to test for this, I separate firms into quartiles each year based on total firm value. I then run the OLS regression specified by 4.1 for each of the quartiles separately. As can be seen in table 4.13, firms in the smallest quartile exhibit this effect to a much higher degree. It is important to remember that all variables are standardized so coefficients indicate the expected standard deviation in leverage for a one standard deviation change in the variable. The t-statistic for the difference between the coefficient for firms in the lowest and largest quartiles is -2.11, indicating that this is stronger for small firms at the 5% level.

Table 4.12: Regression from Equation (4.2) with Post SOX as the Dummy

VARIABLES	(1) Book Leverage	(2) Book Leverage
ln(PPS_stock)	-0.0587** [-2.000]	-0.0729*** [-3.620]
Post SOX Dummy*ln(PPS_stock)	-0.125** [-2.194]	-0.0638** [-2.016]
Post SOX Dummy	0.199** [2.298]	0.241*** [2.679]
ln(Assets)		-0.000144 [-0.00306]
Market to Book		0.0970 [0.905]
PP&E/Assets		0.170*** [4.192]
Mean Industry Leverage Book		0.155*** [3.511]
Abnormal Earnings		-0.0191*** [-2.582]
CEO Ownership		0.0530*** [3.079]
zROA		-0.274*** [-5.356]
Constant	-0.154*** [-5.640]	-0.166*** [-4.613]
Year and Firm Fixed Effects	Yes	Yes
Observations	12,333	12,333
Adj.R-squared	0.594	0.693

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

4.2 Leverage Adjustment Speed

All of the previously discussed results serve to confirm the predictions of my model, however they do not directly speak to the criticism that potentially firms are choosing a low leverage ratio and an executive compensation contract with high PPS due to some underlying omitted variable. Fundamentally, the analysis is missing a causal link. To attempt to establish direct manipulation of leverage ratios by CEOs, I look at the leverage adjustment speed. Since CEOs with high pay performance sensitivity value an underlevered firm, every moment that a firm is underlevered has positive value for the CEO. Therefore, I predict that CEOs with high PPS will manage leverage adjustment actions in order to maximize time in an underlevered state of the world.

Table 4.13: Regression from Equation (4.1) by Size Quartiles with PPS from Stock Only

VARIABLES	(1)	(2)	(3)	(4)
	Book Leverage - Q1	Book Leverage - Q2	Book Leverage - Q3	Book Leverage - Q4
ln(PPS_stock)	-0.285** [-2.498]	-0.0752** [-2.326]	-0.0877** [-2.500]	-0.0391* [-1.794]
ln(Assets)	0.390* [1.908]	0.150 [0.905]	0.0727 [0.535]	-0.239*** [-2.993]
Market to Book	0.435 [1.460]	0.00901 [0.180]	-0.0182 [-0.379]	-0.0286** [-1.996]
PP&E/Assets	0.283*** [4.685]	0.0934 [1.365]	0.125* [1.811]	0.111** [2.327]
Mean Industry Leverage	0.131 [1.508]	0.0897*** [2.671]	0.191*** [5.635]	0.210*** [5.530]
Abnormal Earnings	0.243*** [2.839]	0.510 [1.157]	0.582** [2.137]	0.0471 [0.515]
CEO Ownership	0.171*** [3.192]	0.0617** [2.182]	0.0698* [1.706]	0.0277 [0.776]
ROA	-0.234*** [-8.881]	-0.479*** [-2.976]	-0.173 [-1.200]	-0.315*** [-3.650]
Constant	-0.267 [-1.111]	-0.0525 [-0.381]	-0.0329 [-0.435]	0.485*** [5.242]
Year and Firm Fixed Effects	Yes	Yes	Yes	Yes
Observations	3,076	3,064	3,070	3,061
Adj.R-squared	0.725	0.772	0.794	0.815

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

Table 4.14: Leverage Adjustment Speeds as a Function of PPS

VARIABLES	(1)	(2)	(3)	(4)
	$(L_{i,t+1} - L_{i,t})$	$(L_{i,t+1} - L_{i,t})$	$(L_{i,t+1} - L_{i,t})$	$(L_{i,t+1} - L_{i,t})$
$(L_{i,t+1}^* - L_{i,t})$	0.244** [2.267]			
$\max(L_{i,t+1}^* - L_{i,t}, 0)$		0.451*** [10.33]	0.461*** [10.28]	0.440*** [9.300]
$\min(L_{i,t+1}^* - L_{i,t}, 0)$		0.157 [1.087]	0.153 [1.059]	0.197 [1.350]
High PPS* $\max(L_{i,t+1}^* - L_{i,t}, 0)$			-0.0479** [-2.151]	
High PPS* $\min(L_{i,t+1}^* - L_{i,t}, 0)$			0.108*** [2.759]	
Low PPS* $\max(L_{i,t+1}^* - L_{i,t}, 0)$				0.0420 [0.991]
Low PPS* $\min(L_{i,t+1}^* - L_{i,t}, 0)$				-0.0786* [-1.674]
Constant	0.00731 [1.206]	-0.0208 [-1.187]	-0.0212 [-1.220]	-0.0194 [-1.142]
Year and Firm Fixed Effects	Yes	Yes	Yes	Yes
Observations	10,890	10,951	10,951	10,951
Adj.R-squared	0.199	0.210	0.211	0.211

Robust t-statistics in brackets
 *** p<0.01, ** p<0.05, * p<0.10

I follow Flannery and Rangan 2006 for estimating leverage adjustment speeds. The standard equation for estimating leverage adjustment speeds is

$$(L_{i,t+1} - L_{i,t}) = \lambda(L_{i,t+1}^* - L_{i,t}) + \epsilon_{i,t} \quad (4.3)$$

where $L_{i,t+1}^* = \beta X_{i,t}$, $L_{i,t}$ is the leverage observed for firm i at time t , $L_{i,t}^*$ is the optimal leverage target for firm i at time t , and $X_{i,t}$ is a standard set of controls used for predicting optimal firm leverage (the Frank and Goyal 2008 factors). I estimate a slightly different equation given by

$$(L_{i,t+1} - L_{i,t}) = \lambda_+ \max(L_{i,t+1}^* - L_{i,t}, 0) + \lambda_- \min(L_{i,t+1}^* - L_{i,t}, 0) + \epsilon_{i,t} \quad (4.4)$$

which is identical except I treat the case where target leverage is above current leverage separately from the case in which target leverage is below current leverage. λ_+ is the leverage adjustment speed for when a firm is underlevered relative to its target leverage, and similarly, λ_- is the leverage adjustment speed when a firm is overlevered relative to its target leverage.

Table 4.15: IV Regression - Instrument

VARIABLES	(1) ln(PPS)	(2) ln(PPS_stock)	(3) ln(PPS_option)
CEO Tenure	0.0365*** [5.138]	0.0479*** [6.099]	-0.00165 [-0.268]
ln(Assets)	0.646*** [6.999]	0.332*** [3.408]	0.294*** [4.654]
Market to Book	0.211*** [6.434]	0.137*** [3.610]	0.0687*** [4.775]
PP&E/Assets	-0.0571 [-0.880]	-0.148** [-2.313]	0.0894* [1.702]
Mean Industry Leverage BOOK	-0.0813* [-1.902]	-0.0420 [-1.262]	-0.0386 [-1.542]
Abnormal Earnings	-0.185*** [-2.614]	-0.146 [-1.476]	-0.00823 [-0.121]
CEO Ownership	0.327*** [4.516]	0.361*** [4.620]	-0.125*** [-2.933]
ROA	-0.000685 [-0.0430]	0.00477 [0.391]	-0.00518 [-0.550]
CEO Age	-0.00389 [-0.746]	0.00432 [0.745]	-0.00337 [-0.671]
Firm Age	-0.0518*** [-2.644]	-0.0177 [-1.254]	-0.102*** [-8.074]
Constant	0.158 [0.627]	-0.313 [-1.040]	0.0385 [0.131]
Year and Firm Fixed Effects	Yes	Yes	Yes
Observations	4,587	4,587	4,587
Adj.R-squared	0.681	0.648	0.752

Robust t-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

If a high PPS CEO has an incentive to keep a firm underlevered due to the interaction of bankruptcy risk with PPS, we would expect to see a lower adjustment speed when a firm is underlevered, in order to maximize the time spent in a underlevered state, and when the firm is overlevered, we would expect to see a faster adjustment speed in order to minimize the time spent overlevered. This is exactly the pattern that we see in Table (4.14).

4.3 Instrumental Variable Approach

As discussed previously, it is highly probable that leverage and CEO compensation are jointly determined. If so, the coefficient estimates are likely to be biased. To attack this directly, I use an instrumental variables approach through two stage least

Table 4.16: IV Regression

VARIABLES	(1)	(2)
	Book Leverage	Book Leverage
ln(PPS)	-0.273** [-2.126]	
ln(PPS_stock)		-0.208** [-2.097]
ln(Assets)	0.182** [1.977]	0.0743 [1.144]
Market to Book	0.0371 [1.095]	0.00784 [0.341]
PP&E/Assets	0.288*** [3.619]	0.273*** [3.278]
Mean Industry Book Leverage	0.254*** [5.925]	0.268*** [5.905]
Abnormal Earning	0.191*** [2.658]	0.211*** [2.935]
CEO Ownership	0.105* [1.835]	0.0904* [1.777]
ROA	-0.195*** [-11.07]	-0.194*** [-10.05]
CEO Age	0.00495 [1.517]	0.00691* [1.683]
Firm Age	-0.00629 [-0.595]	0.00418 [0.420]
Year and Firm Fixed Effects	Yes	Yes
Observations	4,493	4,493
Adj.R-squared	.	.

Robust z-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

squares. My two-equation system is given as follows:

$$\ln(\text{PPS}) = \alpha + \beta_0 \text{CEO Tenure} + \beta_{c0} \text{Controls} + \xi \quad (4.5)$$

$$\text{Leverage} = \alpha + \beta_1 \ln(\text{PPS}) + \beta_{c1} \text{Controls} + \epsilon \quad (4.6)$$

where controls are identical as to the basic regression specification, save for the additional controls indicated below.

Identifying an instrument for pay performance sensitivity is not trivial. To my knowledge there are not established instrumental variables for pay performance sensitivity in the existing literature. I use CEO tenure as an instrument for CEO PPS. CEOs tend to accumulate larger equity positions the longer they are with the company as firms have restrictions on executives selling equity. Table 4.15 shows the first stage regression from equation 4.5, in which I regress CEO tenure

on $\ln(\text{PPS})$, $\ln(\text{PPS_stock})$, and $\ln(\text{PPS_option})$, and I find the for both $\ln(\text{PPS})$ and $\ln(\text{PPS_stock})$, CEO tenure is significantly and positively related to PPS. Both statistically and in magnitude, the effect is larger for $\ln(\text{PPS_stock})$, which should be expected, because options have fixed length expirations and should not accumulate with CEO tenure. In accordance with expectation, I observe no relationship between CEO tenure and equity compensation.

CEO tenure is positively related to both firm age and CEO age which may affect firm leverage. Firm age may affect firm access to debt markets, and it has been shown that age is positively related with risk aversion which may affect the CEO decision as to firm leverage. I control for both of these factors in the reported regressions. While I have found nothing in the literature providing a direct link for CEO tenure to firm leverage, Frank and Goyal 2007 do report that CEO tenure is negatively related to firm leverage. However, they do not control for firm age, CEO age and PPS, and if CEO tenure is included in the regression from equation 4.1, CEO tenure is not significantly related. It is my conclusion that Frank and Goyal 2007 were misidentifying the effect. However, their work does not provide an explanation for the relationship, they merely document it.

In table 4.16, I report the results from the two stage least squared regression. Instrumenting for both $\ln(\text{PPS})$ and $\ln(\text{PPS_stock})$, I find identical results to those reported in the basic regression in tables 4.1 and 4.4, with a slightly larger coefficient in both instances.

One potential consideration is survivorship bias. It's possible that a CEO is more likely to maintain his position by reducing leverage, so in unreported results, I restrict my sample to CEOs that have been with the firm for less than five years, in the attempt to mitigate any effect of long tenure CEOs on the level of leverage. Results are qualitatively identical, though slightly less statistically significant, as those reported.

Table 4.17: Regression from Equation (4.1) with Full PPS Measure and CEO Fixed Effects

VARIABLES	(1) Book Leverage	(2) Book Leverage
ln(PPS)	-0.144 [-0.958]	-0.0521 [-0.396]
ln(Assets)		0.843*** [3.794]
Market to Book		-0.00271 [-0.222]
PP&E/Assets		0.191 [1.126]
Mean Industry Book Leverage		0.182*** [2.691]
Abnormal Earnings		2.365*** [4.871]
CEO Ownership		0.950 [1.636]
ROA		0.196 [0.878]
Constant	0.278* [1.756]	0.0325 [0.117]
Year, Firm, and CEO Fixed Effects	Yes	Yes
Observations	225	225
Adj.R-squared	0.883	0.909

Robust t-statistics in brackets
 *** p<0.01, ** p<0.05, * p<0.10

4.4 Further Discussion

I've looked at a number of other potentially relevant variables to rule out other explanations or strengthen my interpretation that I don't report in this draft, but I would like to discuss them briefly. I use volatility as a measure of likelihood of default, but this is a very indirect measure of default probability. A natural measure to use is the Altman z-score, which measures the financial health of a firm directly. I have used the Altman z-score in place of volatility and the results are consistent, though not significant. However, the Altman z-score indicates firms that are already in financial distress. Once a firm enters financial distress, the CEO is likely to already be in the process of removal, so it's unclear what the relationship between PPS and leverage means for a financially distressed firm.

Further, the measure I use for PPS is not scaled by total compensation, following

the standard PPS literature. This is consistent with the notion of CARA utility that I include in my model. However, for an agent with an alternative specification of utility, the scaled measure may make more sense. The difficulty with scaling PPS is determining the correct scaling factor since PPS is potentially an accumulation of many years of stock and option grants. I did scale PPS by total salary and bonus for the given year and found identical results, but it's not clear that this is an appropriate scaling factor. I will provide those regressions upon request.

I use a 20% threshold for all regressions for equation (4.2), but the results are robust to other thresholds as well. I chose 20% to maintain a large enough number of observations captured by the dummy variable, but I've looked at 5%, 10%, 30%, and 40% and the results are consistent.

A potential criticism of my explanation is that both firm leverage and executive compensation is endogenous. To examine this, I have included executive fixed effects to my current regressions. All results are essentially unchanged since executive mobility is relatively low, so firm fixed effects already control for most of the variation. To further explore this, I restrict my sample to CEOs that have been employed by at least two of the firms in the sample, and I then estimate equation (4.1) including CEO fixed effects. This severely restricts my sample (there are only 40 executives that have worked at two or more firms in my sample), so the results are no longer significant, but the sign is still consistent, as reported in Table (4.17). I am currently working to enlarge this sample using a technique described in Graham, Li, and Qiu 2011, but I am not able to report those results at this time.

One element of compensation that I do not include is severance payments. If termination in bankruptcy is the primary concern for the executive, a natural way to extend the compensation contract is to ensure a large positive payment to the executive in the case of termination due to default. There are several difficulties with studying severance payments that make them impractical for the current iteration of

this line of research. First, data is not widely reported for severance payments that are part of an explicit agreement. Second, a large proportion of severance agreements are implicit in nature, and therefore unobservable. Thirdly and most directly related to my work, Berk, Stanton, and Zechner (2010) argue that it is unlikely that firms have the legal ability to commit to a severance payment in bankruptcy. If the severance payment is not relevant to the state of the world in which default occurs, it should have no effect on my results. Further, if severance payments are made in default, the omission from my specification should only serve to weaken my results. However, I am currently attempting to acquire data on severance payments to test the effect.

Conclusion

I am able to establish the robustness of the negative relationship between pay performance sensitivity and leverage for the first time. This result is in contradiction to the standard result that firms are underlevered and that greater PPS leads the CEO to strictly improve firm performance. My results are not in contradiction to work that claims that firm value increases as PPS increases, since I'm only examining a single choice out of the sets of choices that the typical executive makes. Further, I don't rely on the necessity of widespread firm underleverage for my results to be meaningful, for the result is still novel and significant. Nor does my interpretation of the results depend on underleverage.

I develop and partially solve a model that predicts the empirically observed negative relationship and makes further predictions. I show that the relationship becomes more negative for high volatility firms, and more negative for firms that are unlikely to take disciplinary actions against the current CEO. Both of these results are entirely new to the best of my knowledge. I further demonstrate the robustness of my results by examining a restricted measure of pay performance sensitivity. I also show that this affect varies by firm size with small firms, the firms more likely to

experience bankruptcy, having a significantly stronger relationship between PPS and leverage. Since it is likely that personal default costs increase post-SOX, I test to see if the relationship strengthens, and it does become significantly stronger post-SOX.

This effect seems to be primarily due to the stock holdings of CEOs and not the option holdings. This is inconsistent with the agency cost of debt being the primary factor affecting this relationship. Further, in my instrumental variable approach, I use an instrument that is unrelated to option PPS, and I still observe a significant relationship. To the best of my knowledge, there is not another explanation that can account for the observed empirical relationship.

I am extending the analysis currently present in this line of research by examining heterogeneity in personal costs of bankruptcy following some of the analysis in Eckbo, Thorburn, and Wang 2012. While initial results are promising, there is more work to be done to fully examine this heterogeneity. Further, I am working to exploit events where firms merge or split to enable me to control for corporate governance. Finally, There is heterogeneity in the extent to which CEOs can legally hedge their own exposure to equity risk in their own firm (Gao 2010), and I am currently gathering data to exploit this heterogeneity.

Endogeneity of the compensation contract is a significant concern. It's entirely possible that CEOs are selecting firms that would optimally choose lower leverage and provide greater pay performance sensitivity to their executives. I provide evidence that this isn't a complete story because I demonstrate that leverage adjustment speeds are affected by PPS in a way consistent with CEOs trying to actively manage debt downwards. Further, I provide direct evidence through instrumenting for CEO PPS using CEO tenure. I show that with CEO tenure instrumenting for PPS, results are identical.

Though the model is currently quite simplistic, it is not trivial to solve. I am currently exploring other modeling options to increase the analytical richness of my

solution. In addition, I am exploring the structure of optimal compensation contracts from the perspective of the firm. It may be that firms are structuring contracts optimally, and that providing incentives for the CEOs to increase leverage is too costly. However, I suspect that this is an area of unnecessary wealth appropriation by CEOs.

The application of this dissertation is that boards should potentially consider implementing direct incentives for the executive to increase leverage. However, if there is an information asymmetry between firms governing body and the typical executive, it may be prudent for firms to regularly have an independent auditor examine the leverage choices of the firm and take disciplinary action against the current CEO in the case of a sub-optimal choices.

Going forward, it seems fruitful to continue examining other unexpected, and, therefore, likely unintended, consequences of the current structure of executive compensation. In this dissertation, I show that the interaction of the standard compensation contract with the well documented cost of default borne by executives can have surprising and financially significant results. Whether this interaction drives other firm decisions that may affect bankruptcy risk is still an open question.

Appendix A

Proofs of Theoretical Results

Proof of Theorem 2. From equation 2.2, we have that the agent is trying to maximize

$$U(x, L) = \gamma(w_0 + \alpha(1 + tL)x + \phi L - \frac{kx^2}{2} - \frac{1}{2}\gamma\alpha^2\sigma^2) - \ln(e^{d\gamma}\Phi(\epsilon'_D) + (1 - \Phi(\epsilon'_D))).$$

Therefore, the first order conditions state that

$$\frac{\partial U}{\partial x} = \gamma\alpha(1 + tL) - \gamma kx - \frac{(e^{d\gamma} - 1)\Phi'(\epsilon'_D)}{(e^{d\gamma} - 1)\Phi(\epsilon'_D) + 1} \left(\frac{(1 - t)L - 1}{\sigma} \right) = 0 \quad (\text{A.1})$$

and

$$\frac{\partial U}{\partial L} = \gamma\alpha tx + \gamma\phi - \frac{(e^{d\gamma} - 1)\Phi'(\epsilon'_D)}{(e^{d\gamma} - 1)\Phi(\epsilon'_D) + 1} \left(\frac{(1 - t)x}{\sigma} \right) = 0 \quad (\text{A.2})$$

Equations A.1 and A.2 can be jointly solved for x , where

$$x = \frac{\alpha + \sqrt{\alpha^2 - 4k\phi((1 - t)^2L - (1 - t))}}{2k(1 - t)} \quad (\text{A.3})$$

From this point onward in the proof, I will use the assumption that the principal cannot or does not contract on leverage, i.e. $\phi = 0$. Then equation A.1 becomes

$$x = \frac{\alpha}{k(1 - t)} \quad (\text{A.4})$$

This implies that ϵ'_D , given by equation 2.4, can be written as

$$\epsilon'_D = \frac{\alpha(((1-t)L - 1)^{\frac{1}{(1-t)k}} + \gamma\sigma^2)}{\sigma} \quad (\text{A.5})$$

Notice that from equation A.5 and assumption 1, $\epsilon'_D > 0$ and $\frac{\partial \epsilon'_D}{\partial L} > 0$. If I further assume that d is arbitrarily large, then equation A.2 converges to

$$\frac{\partial U}{\partial L} = \gamma\alpha t - \frac{\Phi'(\epsilon'_D)}{\Phi(\epsilon'_D)} \left(\frac{(1-t)}{\sigma} \right) = 0 \quad (\text{A.6})$$

or

$$\frac{\gamma\alpha\sigma t}{1-t} = \frac{\Phi'(\epsilon'_D)}{\Phi(\epsilon'_D)}. \quad (\text{A.7})$$

This can be totally differentiated to give the following

$$\frac{dL}{d\alpha} = \frac{\gamma k t \sigma^2}{(1-t)^2} \frac{2(\Phi(\epsilon'_D))^2}{2\Phi(\epsilon'_D)\Phi''(\epsilon'_D) - (\Phi'(\epsilon'_D))^2} - \gamma\sigma \quad (\text{A.8})$$

where $\Phi(\cdot)$ is the CDF of the normal distribution. Using Properties of normal distributions, it can be shown that $\frac{2(\Phi(\cdot))^2}{2\Phi(\cdot)\Phi''(\cdot) - (\Phi'(\cdot))^2} < 0$ for any input greater than or equal to zero. Since by assumption of a sufficiently large k and $L \geq 0$, equation A.5 shows that $\epsilon'_D > 0$. Therefore, $\frac{dL}{d\alpha} < 0$. □

Proof of Theorem 3. Equation A.8 says

$$\frac{dL}{d\alpha} = \frac{\gamma k t \sigma^2}{(1-t)^2} \frac{2(\Phi(\epsilon'_D))^2}{2\Phi(\epsilon'_D)\Phi''(\epsilon'_D) - (\Phi'(\epsilon'_D))^2} - \gamma\sigma$$

and it can be further shown that $\frac{2\Phi(\cdot)\Phi''(\cdot) - (\Phi'(\cdot))^2}{2(\Phi(\cdot))^2}$ approaches 0 asymptotically from below, using the properties of normal distributions, for all series starting from greater than zero and approaching positive infinity. Therefore, as σ is increasing, ϵ'_D is increasing. Therefore, $\frac{dL}{d\alpha}$ becomes more negative. So, $\frac{d^2L}{d\alpha d\sigma}$. □

Proof of Theorem 4. As ϕ increases, the relative trade off between the personal benefits of debt (the share in the tax shield and the personal compensation for increased leverage) becomes relatively larger, while the personal costs of default stays constant. Therefore, a small change in the exposure to firm equity, an increase in α , affects the risk adjusted probability of default identically to the previous case, but the relative benefits of leverage are still larger. This leads to a smaller change in leverage, i.e.

$$\frac{d^2L}{d\phi d\alpha} > 0.$$

□

Appendix B

Robustness Checks

In this appendix, I include additional tables and figures that are relevant for robustness checks. Specifically, figures B.1 and B.2, show a graphical univariate analysis of the relationship between the CEOs equity value and the chosen firm leverage relative to the other firms in the industry. I demean the leverage on an annual basis relative to the two digit SIC industry for the firm and chart the demeaned leverage against the decile of the value of the CEOs equity holding. The relationship observed in figures 1.1 and 1.2 are nearly identical in trend to figures B.1 and B.2. This indicates graphically that the observed relationship is not driven by industry specific factors for leverage.

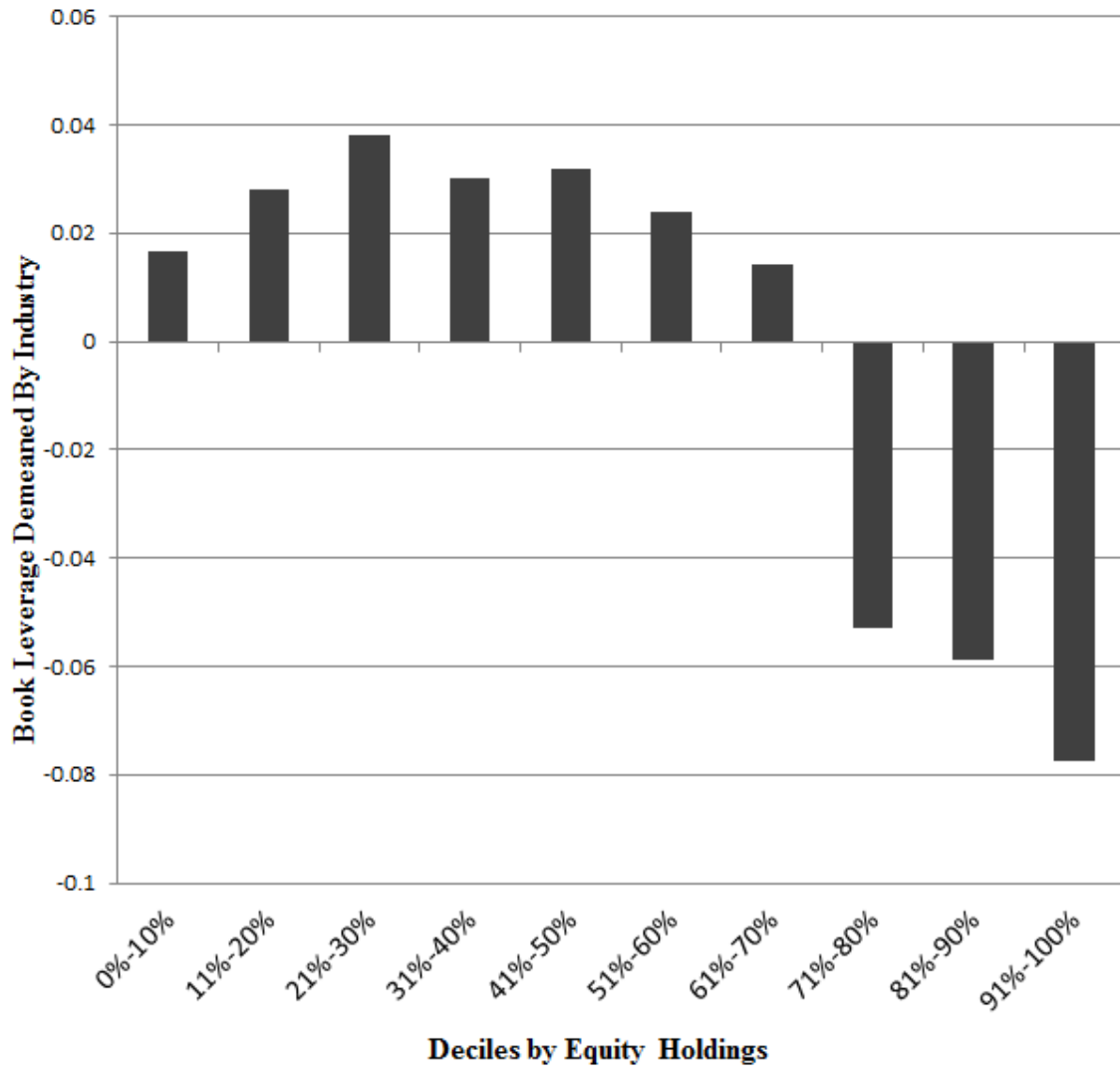


FIGURE B.1: Industry Demeaned Book Leverage by Deciles of Pay Performance Sensitivity

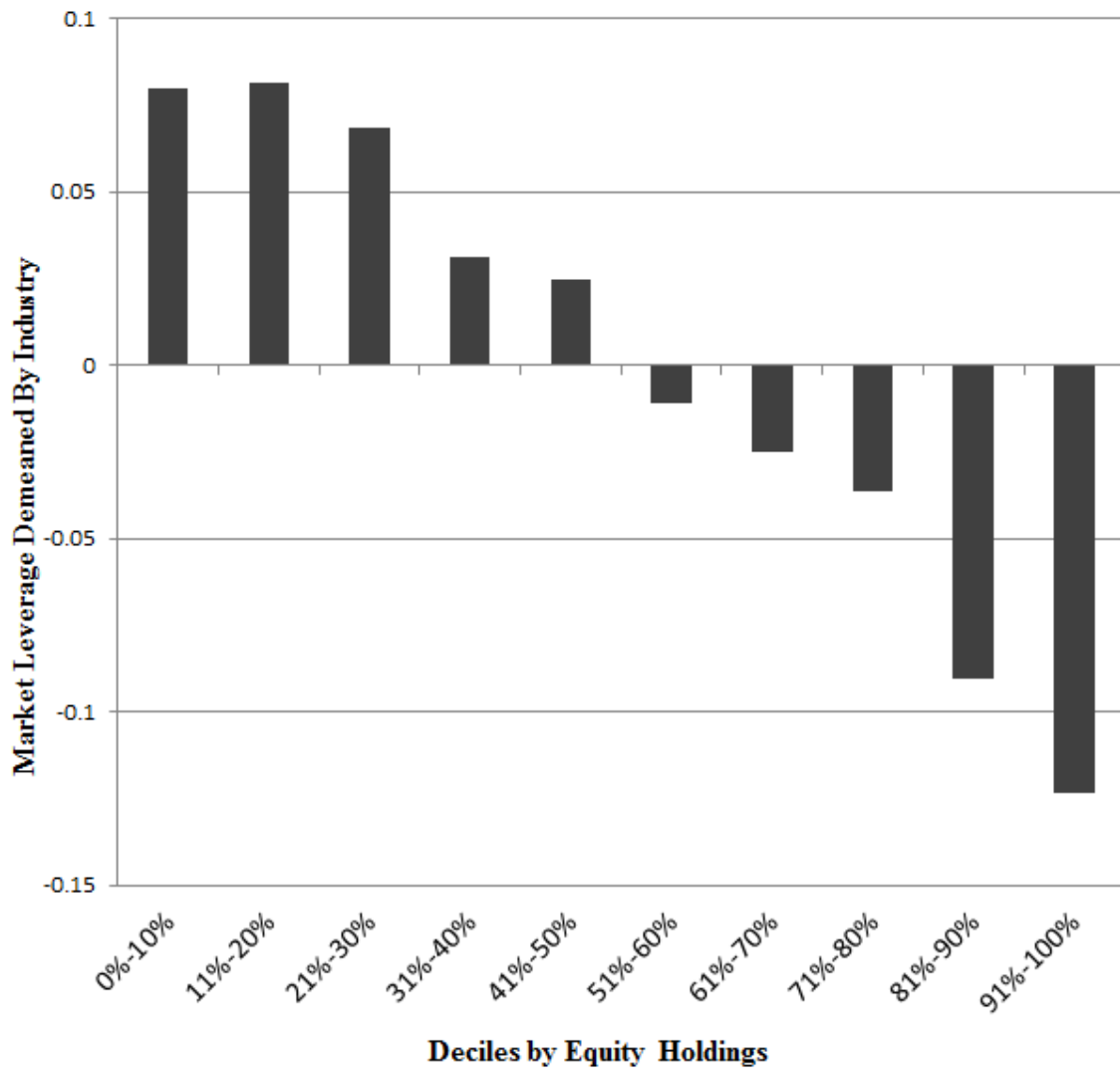


FIGURE B.2: Industry Demeaned Market Leverage by Deciles of Pay Performance Sensitivity

Appendix C

A Natural Experiment

C.1 Introduction

The main difficulty with examining the effect of executive compensation on firm financing decisions is that compensation and firm financing decisions are both set endogenously by the firm. It is an open question as to how much the members of the CEO compensation committee consider the effect of the compensation contract on firm leverage decision. In this dissertation, I attempt to remove some of the endogeneity inherent in the question by using CEO tenure as an instrument for the CEO equity stake. However, a natural experiment due to a legal change is a cleaner and clearer method to examine this question.

In order to try to address this endogeneity, I will use the change in the tax law due to Internal Revenue Code (IRC) 162(m), which went into affect January 1st, 1994. In IRC 162(m), the amount that a firm can expense due to CEO compensation changed. Before IRC 162(m), a firm could expense the entirety of CEO compensation, but afterwards, the firm could only expense up to \$1 million. However, there is an exception in the law for performance based compensation; the entirety of performance

based compensation can be expensed. Therefore, after the law took effect, for firms that paid a salary at or above \$1 million, there was a significant tax advantage of changing the composition of the compensation contract to more performance based compensation and less salary compensation.

However, firms did not shift compensation equally to all forms of performance based compensation. For various reasons, Perry and Zenner (2001) report that the majority of the increase in performance based compensation came from an increase in option grants. The results that I report in chapter 4 indicate that the important consideration for CEOs when they are deciding leverage ratios is the size of their stock grants, not the size of the option grants. Therefore, this natural experiment is not directly applicable to confirming the results reported in chapter 4, but it can be used to confirm that option grants do not have a significant negative relationship with firm leverage decisions.

The natural experiment has a fairly clear interpretation. The treatment group are firms that pay the CEO close to \$1 million in salary. If option grants do significantly affect the leverage ratio, we should see a marked decrease in the chosen leverage ratio for firms these firms post 1994 relative to firms that pay the CEO less. However, the firms in the two groups are not identical. As may be expected, firms that pay the CEO close to \$1 million in salary are also firms that are on average significantly larger than firms that pay the CEO much less, so it is important to carefully control for firm size and other factors that may co-vary with firm size.

This dissertation is not the first work to use this natural experiment to look at the relationship between option grants and leverage ratios. Tosun (2012) looks at the effect of within year option grants on the leverage ratio. His methodology differs significantly than mine over several measures. First, he uses a within year measure of option grants, and I use a measure of the entire option holdings of the CEO. Second, his main results rely on a misspecified model and do not have any direct

bearing on the issue of whether or not option grants affect leverage ratios, whereas I rely strictly on a standard diff-in-diff framework and instrumental variable approach. Third, in his correctly specified instrumental variable approach that he includes as a robustness test, he finds only marginal significance. However, he does not cluster standard errors, though he has a panel data set, and his results lose all significance once standard errors are correctly clustered.

Further, the explanation that Tosun (2012) gives for the negative relationship that he marginally reports is different, though not contradictory, to mine. He relies on the agency theory of debt. Specifically, since option grants give a risk neutral, though not necessarily a risk averse, CEO an incentive to increase firm volatility, the CEO is more likely to take on riskier projects. Since the CEO is going to take on riskier projects, debtors understand this, and they bear a disproportionate level of the risk from these riskier projects. The debtors then raise interest rates on debt, which causes the firm to reduce leverage.

This agency cost of debt explanation relies fundamentally on the convexity of option payoffs inducing risky behaviour by the CEO. If the CEO has no incentive to take on riskier projects, the firms leverage ratio should not change. However, if a CEO is risk averse, an increase in the option grants has an ambiguous effect on risk taking (see Ross (2004) and Lambert et al. (1991)). My results suggest that the fundamental concern for the CEO is in actively managing the firm leverage to manage personal risk due to default.

What I find from this natural experiment is that option grants do not significantly affect leverage. This is consistent with the previous empirical results. However, I also demonstrate that this natural experiment is not valid for stock grants. I see a marginal decrease in the leverage ratio due to stock grants if I use this natural experiment as an instrumental variable for stock grants, however it is not significant. Since this is a poor instrument for stock grants, I do not interpret these results as

being inconsistent with results reported elsewhere.

C.2 Results from a Natural Experiment

I use identical data sources and variable definitions as reported in chapter 3. However, I restrict my sample to the years 1993-1997. I only include firms that have all variables of interest over the entire sample period. This leaves me with 385 unique firms and 1920 firm year observations. This is the sample window that Tosun (2012) uses, and it is the sample window that seems to be the most conducive to positive results consistent with the agency cost of debt. I have examined many other sample windows as robustness tests, and results are generally weaker for other sample windows.

I define the treatment group as all firms that pay the CEO a salary greater than or equal to \$.9 million. I have used various salary cutoffs and results are strictly weaker. The post variable is a dummy variable that takes the value 1 if the year is greater than or equal to 1994. Therefore, the only year in my sample that is pre-effect is 1993.

I do include one additional variable as a test. The agency cost of debt theory that Tosun (2012) provides as an explanation for his results relies on the sensitivity of option grants to the underlying volatility of the stock price. Therefore, I include the option vega as a measure of volatility sensitivity. The option vega is calculated as the variable VOLSEN used in Brockman et al. (2010), and it is a measure of the change in option value due to the change in the underlying stock volatility. This is the vega of the entire portfolio of options that the CEO holds, so this should be the measure with the most explanatory power for the agency cost of debt.

The first test that I perform is a univariate test of the change in the means of the variables of interest between the years 1993 and 1997 for both the treated and untreated group. Results are reported in table C.1. First, there is no significant

Table C.1: Change in Mean Values Pre and Post Affect

	Untreated		Treated	
	Pre - 1993	Post - 1997	Pre - 1993	Post - 1997
PPS_stock (in millions)	45.2	57.4	106	217
Difference		12.2		111
T-stat		[.7297]		[1.12]
BS Option Value (in millions)	.497	1.26	1.27	5.77
Difference		.763***		4.50***
T-stat		[3.2]		[3.12]
Vega (in thousands)	35.2	77.5	123	315
Difference		42.3***		192***
T-stat		[3.21]		[3.44]
PPS_options (in thousands)	74.7	210	196	856
Difference		135***		660***
T-stat		[3.23]		[5.32]
Book Leverage	.460	.490	.569	.563
Difference		.030**		-.006
T-stat		[1.931]		[-.215]
Number of Firms	343	306	41	78

change in stock grants for either the treated or untreated group. The treated firms are significantly larger, so the stock grants, and therefore the change, is much larger for the treated group, but neither change is statistically significant. However, we do see significant changes for all of the measures of the option grants. For both the treated and the untreated group we see option grants increase on all measures in a statistically significant fashion. However, the magnitude of the increase is much larger for the treated group than the untreated. Further, the leverage seems to vary in a manner consistent with increased option grants reducing leverage ratios. Specifically, leverage increases significantly for the control group, but decreases, albeit insignificantly, for the treatment group.

Table C.1 seems to indicate that this is a useful natural experiment for examining the effect of option grants on leverage, though it also indicates that this natural experiment is unlikely to be useful for examining the effects of stock grants. In addition, it suggests, that there may be some effect given the differential changes in leverage for the treated and untreated groups.

Table C.2: Diff-in-Diff from a Natural Experiment

VARIABLES	(1) Book Leverage	(2) Book Leverage
Post	0.0500** [2.273]	-0.0171 [-0.762]
Treated	0.453*** [5.476]	-0.0790 [-1.061]
Post*Treated	-0.0353 [-0.876]	0.0276 [0.754]
ln(Assets)		0.357*** [10.18]
Market to Book		-0.158*** [-3.531]
PP&E/Assets		-0.0474 [-1.538]
Mean Industry Leverage BOOK		0.200*** [4.359]
Abnormal Earnings		2.891*** [5.183]
CEO Ownership		-0.0704** [-2.184]
ROA		-0.318*** [-2.622]
Constant	-0.0761 [-1.579]	0.0442 [1.118]
Observations	1,920	1,920

Robust t-statistics in brackets
 *** p<0.01, ** p<0.05, * p<0.10

To move beyond the univariate analysis, I use a standard diff-in-diff specification with clustered standard errors. Results are reported in table C.2. For both the specifications with and without controls, there is no significant effect for the post*treated variable, and for the specifications with controls, I find a coefficient that is positive. This indicates that, using a diff-in-diff approach with this natural experiment, there is no evidence that option grants affect leverage ratios.

A diff-in-diff approach is not necessarily ideal for determining the effect of various components of the compensation contract on firm leverage, due to the inability to disentangle the effects of the increase, though insignificant, of stock grants versus option grants. In order to more directly examine the effect of the components of the compensation contract considered separately, I use an instrumental variable approach. I use the post, treated, and post*treated variables as instruments for both stock grants and the various measures of option grants that I use. In order to test

Table C.3: IV Regression - Instrument for the Natural Experiment

VARIABLES	(1) ln(PPS_stock)	(2) ln(PPS_option)	(3) ln(BS Option Value)	(4) ln(Vega)
post	0.00814 [0.285]	0.138*** [4.258]	0.107** [2.416]	0.103*** [3.141]
Treated	0.132 [1.021]	-0.0790 [-0.477]	-0.340** [-2.248]	-0.0575 [-0.339]
Post*Treated	0.0697 [0.589]	0.281** [2.317]	0.222 [1.533]	0.244* [1.937]
ln(Assets)	0.225*** [6.012]	0.248*** [5.724]	0.225*** [6.127]	0.250*** [5.853]
Market to Book	0.216*** [5.213]	0.0666 [0.850]	0.0469 [0.836]	-0.0321 [-0.442]
PP&E/Assets	-0.0359 [-1.176]	-0.0887** [-2.302]	-0.0441 [-1.562]	-0.0739* [-1.920]
Mean Industry Leverage BOOK	-0.00453 [-0.117]	-0.0525 [-1.015]	-0.0692 [-1.627]	-0.0541 [-1.079]
Abnormal Earnings	-0.801 [-1.625]	-0.0552 [-0.153]	-0.426 [-0.581]	0.0935 [0.252]
CEO Ownership	0.452*** [9.699]	-0.321*** [-5.725]	-0.234*** [-7.211]	-0.306*** [-5.618]
ROA	0.352*** [4.346]	-0.0881 [-0.785]	-0.110 [-1.268]	-0.123 [-1.115]
Constant	0.00819 [0.194]	-0.245*** [-4.809]	-0.0304 [-0.613]	-0.243*** [-4.887]
Observations	1,920	1,920	1,920	1,920

Robust z-statistics in brackets
 *** p<0.01, ** p<0.05, * p<0.10

the validity of these as an instrument, I report the first stage of the instrumental variable approach in table C.3.

As can be seen from table C.3, the natural experiment seems to be a reasonably good specification for the various option measures in that in all specifications at least two of the variables from the set post, treated, and post*treated do significantly effect the option measure. However, for stock grants, the effect is always insignificant. I conclude from this that this natural experiment is a good instrument for option grants but not for stock grants.

I use a standard instrumental variable approach where I include the instrumented variables from both stock and option grants and all variations on the option grant metric. Tables C.4, C.5, and C.6 report the results.

As can be seen in tables C.4, C.5, and C.6, the instrumental variable approach does not support option grants significantly affecting the leverage choice. This is

Table C.4: IV Regression for the Natural Experiment with PPS

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage
ln(PPS_stock)	-0.505 [-0.784]		-0.592 [-0.707]
ln(PPS_option)		-0.156 [-0.825]	0.0457 [0.277]
ln(Assets)	0.475*** [2.719]	0.385*** [6.289]	0.485** [2.458]
Market to Book	-0.0458 [-0.298]	-0.148*** [-3.132]	-0.0294 [-0.157]
PP&E/Assets	-0.0676 [-1.628]	-0.0600* [-1.735]	-0.0671 [-1.566]
Mean Industry Leverage BOOK	0.197*** [3.940]	0.192*** [4.018]	0.199*** [3.790]
Abnormal Earnings	2.477*** [3.401]	2.864*** [5.216]	2.414*** [2.916]
CEO Ownership	0.158 [0.532]	-0.122* [-1.794]	0.213 [0.505]
ROA	-0.137 [-0.533]	-0.329*** [-2.776]	-0.102 [-0.308]
Constant	0.0400 [0.826]	0.00118 [0.0312]	0.0487 [0.727]
Observations	1,920	1,920	1,920

Robust z-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

especially true if one is arguing that increased option grants decreases the leverage ratio because in all specifications once stock grants are included the estimated coefficient on option grants is positive, and in table C.5, it is always positive.

However, tables C.4, C.5, and C.6 do estimate a negative coefficient, though always insignificant, for the effect of stock grants on leverage decisions. This is consistent with the direction of the effect estimated elsewhere in the dissertation, but the coefficients are larger for this IV approach and insignificant. However, as stated previously, this natural experiment is not appropriate for stock grants, so I would not expect to see statistical significance for the instrumented stock grants.

C.3 Conclusion to a Natural Experimental

These results only weakly confirm my main results. This natural experiment does not find any significance for either stock or option grants affecting firm leverage.

Table C.5: IV Regression for the Natural Experiment with Black-Scholes Value for Options

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage
ln(PPS_stock)	-0.505 [-0.784]		-0.498 [-0.803]
ln(BS Option Value)		0.0967 [0.435]	0.0230 [0.125]
ln(Assets)	0.475*** [2.719]	0.322*** [5.778]	0.469*** [2.932]
Market to Book	-0.0458 [-0.298]	-0.163*** [-3.528]	-0.0482 [-0.330]
PP&E/Assets	-0.0676 [-1.628]	-0.0413 [-1.271]	-0.0664* [-1.686]
Mean Industry Leverage BOOK	0.197*** [3.940]	0.209*** [4.198]	0.199*** [3.861]
Abnormal Earnings	2.477*** [3.401]	2.941*** [4.936]	2.494*** [3.503]
CEO Ownership	0.158 [0.532]	-0.0487 [-0.740]	0.161 [0.524]
ROA	-0.137 [-0.533]	-0.304** [-2.415]	-0.137 [-0.530]
Constant	0.0400 [0.826]	0.0142 [0.435]	0.0389 [0.846]
Observations	1,920	1,920	1,920

Robust z-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

However, the results are consistent with my main findings. Specifically, I find that the negative relationship between pay performance sensitivity and firm leverage is primarily driven by the stock component of pay performance sensitivity elsewhere in this dissertation, and here I find that, though the natural experiment seems to be valid for option grants, there seems to be no significant causal negative relationship between option compensation and firm leverage choice.

I also find no significance when I examine stock grants, but this natural experiment is not appropriate for stock grants since IRC 162(m) has been shown to cause firms to shift from salary and primarily towards option grants and not stock grants. Therefore, a null result does not diminish my findings elsewhere because one would not expect to find significance when the natural experiment is inappropriate.

This does help me to distinguish between potential competing explanations for the negative relationship between pay performance sensitivity and leverage. Specifically,

Table C.6: IV Regression for the Natural Experiment with Vega

VARIABLES	(1) Book Leverage	(2) Book Leverage	(3) Book Leverage
ln(PPS_stock)	-0.505 [-0.784]		-0.634 [-0.652]
ln(Vega)		-0.194 [-0.813]	0.0654 [0.262]
ln(Assets)	0.475*** [2.719]	0.396*** [5.427]	0.491** [2.306]
Market to Book	-0.0458 [-0.298]	-0.165*** [-3.611]	-0.0149 [-0.0652]
PP&E/Assets	-0.0676 [-1.628]	-0.0607* [-1.703]	-0.0680 [-1.527]
Mean Industry Leverage BOOK	0.197*** [3.940]	0.189*** [3.906]	0.200*** [3.681]
Abnormal Earnings	2.477*** [3.401]	2.891*** [5.245]	2.371** [2.553]
CEO Ownership	0.158 [0.532]	-0.131* [-1.667]	0.237 [0.469]
ROA	-0.137 [-0.533]	-0.339*** [-2.781]	-0.0837 [-0.213]
Constant	0.0400 [0.826]	-0.00797 [-0.177]	0.0543 [0.636]
Observations	1,920	1,920	1,920

Robust z-statistics in brackets
*** p<0.01, ** p<0.05, * p<0.10

the results reported here suggest that the agency cost of debt working through the distorted incentives due to option compensation does not strongly contribute to the firm leverage decision. Given that these results are consistent with my other empirical work that demonstrates that stock grants are the main driver, it weakly supports the theory proposed in this dissertation that the interaction of the CEO personal wealth portfolio and labor risk in default is the main driver of the negative relationship.

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Biography

The author of this dissertation is Michael Albert. He was born at a young age in Harrisonburg, VA on the day of July 22nd of the year of our Lord 1986. Michael was a precocious youngster, constantly getting into trouble, yet he loved to read, so when he was sent to timeout, he could just read a book, which might have been a tactical error on the part of his parents. This tendency towards the academic led him to gravitate towards academics, as might be expected.

Gravitate is a fine word to use, because he fell in love with the study of physics in no small part due to Einstein's theory on gravity, to the detriment of other pursuits. This led him to enroll in college at the age of 15 to continue his affair with physics. At this place of higher education, James Madison University, he fell in with a bad crowd and studied mathematics and quantitative finance as well, leading him to graduate with a Bachelor of Science in quantitative finance and a Bachelor of Science in mathematics with a heavy minor in physics.

This focus on mathematics and quantitative finance was not as marketable as he hoped it would be, so he enrolled at Duke University in the PhD in Business Administration with a focus on finance. There he developed an interest in computer science, and decided it would be dandy if he had a Masters of Computer Science as well. In the year of our Lord 2013, he graduated with both degrees, and went on to have many fine and varied adventures.